

GENERAL SMILE ASYMPTOTICS WITH BOUNDED MATURITY

FRANCESCO CARAVENNA AND JACOPO CORBETTA

ABSTRACT. We provide explicit conditions on the distribution of risk-neutral log-returns which yield sharp asymptotic estimates on the implied volatility smile. We allow for a variety of asymptotic regimes, including both small maturity (with arbitrary strike) and extreme strike (with arbitrary bounded maturity), extending previous work of Benaim and Friz [BF09]. We present applications to popular models, including Carr-Wu finite moment logstable model, Merton's jump diffusion model and Heston's model.

1. INTRODUCTION

The price of a European option is typically expressed in terms of the Black&Scholes *implied volatility* $\sigma_{\text{imp}}(\kappa, t)$ (where κ denotes the log-strike and t the maturity), cf. [G11]. Since exact formulas for a given model are typically out of reach, an active line of research is devoted to finding asymptotic expansions for $\sigma_{\text{imp}}(\kappa, t)$, which can be useful in many respects, e.g. for a fast calibration of some parameters of the model. Explicit asymptotic formulas for $\sigma_{\text{imp}}(\kappa, t)$ also allow to understand how the parameters affect key features of the volatility surface, such as its slope, and what are the possible shapes that can actually be obtained for a given model. Let us mention the celebrated Lee moment's formula [L04] and more recent results [BF08, BF09, T09, G10, FF12, MT12, GL14, FJ09, RR09, GMZ14].

A key problem is to link the implied volatility *explicitly* to the distribution of the risk-neutral log-return X_t , because the latter can be computed or estimated for many models. The results of Benaim and Friz [BF09] are particularly appealing, because they connect directly the asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ to the *tail probabilities*

$$\bar{F}_t(\kappa) := \mathbb{P}(X_t > \kappa), \quad F_t(-\kappa) := \mathbb{P}(X_t \leq -\kappa). \quad (1.1)$$

Their results, which are limited to the special regime of extreme strike $\kappa \rightarrow \pm\infty$ with fixed maturity $t > 0$, are based on the key notion of *regular variation*, which is appropriate when one considers a single random variable X_t (since t is fixed). This leaves out many interesting regimes, notably the much studied case of small maturity $t \rightarrow 0$ with fixed strike κ .

In this paper we provide a substantial extension of [BF09]: we formulate a suitable generalization of the regular variation assumption on $\bar{F}_t(\kappa)$, $F_t(\kappa)$ which, coupled to suitable moment conditions, yields the asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ in essentially *any regime of small maturity and/or extreme strike* (with bounded maturity). We thus provide a unified approach, which includes as special cases both the regime of extreme strike $\kappa \rightarrow \pm\infty$ with fixed maturity $t > 0$, and that of small maturity $t \rightarrow 0$ with fixed strike κ . Mixed regimes, where κ and t vary simultaneously, are also allowed. This flexibility yields asymptotic formulas for the volatility surface $\sigma_{\text{imp}}(\kappa, t)$ in open regions of the plane.

Date: July 7, 2016.

2010 *Mathematics Subject Classification.* Primary: 91G20; Secondary: 91B25, 60G44.

Key words and phrases. Implied Volatility, Asymptotics, Volatility Smile, Tail Probability.

In Section 3 we illustrate our results through applications to popular models, such as Carr-Wu finite moment logstable model and Merton’s jump diffusion model. We also discuss Heston’s model, cf. §3.3. In a separate paper [CC15] we consider a stochastic volatility model which exhibits multiscaling of moments, introduced in [ACDP12].

The key point in our analysis is to connect explicitly the asymptotic behavior of the tail probabilities $\bar{F}_t(\kappa), F_t(\kappa)$ to call and put option prices $c(\kappa, t), p(\kappa, t)$ (cf. Theorems 2.3, 2.4 and 2.7). In fact, once the asymptotics of $c(\kappa, t), p(\kappa, t)$ are known, the behavior of the implied volatility $\sigma_{\text{imp}}(\kappa, t)$ can be deduced in a model independent way, as recently shown Gao and Lee [GL14]. We summarize their results in §2.4 (see Theorem 2.9), where we also give an extension to a special regime, that was left out from their analysis (cf. also [MT12]).

The paper is structured as follows.

- In Section 2 we set some notation and we state our main results.
- In Section 3 we apply our results to some popular models.
- In Section 4 we prove Theorem 2.9, linking option price and implied volatility.
- In Section 5 we prove our main results (Theorems 2.3, 2.4 and 2.7).
- Finally, a few technical points have been deferred to the Appendix A.

2. MAIN RESULTS

2.1. The setting. We consider a generic stochastic process $(X_t)_{t \geq 0}$ representing the log-price of an asset, normalized by $X_0 := 0$. We work under the risk-neutral measure, that is (assuming zero interest rate) the price process $(S_t := e^{X_t})_{t \geq 0}$ is a martingale. European call and put options, with maturity $t > 0$ and a log-strike $\kappa \in \mathbb{R}$, are priced respectively

$$c(\kappa, t) = \mathbb{E}[(e^{X_t} - e^\kappa)^+], \quad p(\kappa, t) = \mathbb{E}[(e^\kappa - e^{X_t})^+], \quad (2.1)$$

and are linked by the *call-put parity* relation:

$$c(\kappa, t) - p(\kappa, t) = 1 - e^\kappa. \quad (2.2)$$

As in [GL14], in our results *we take limits along an arbitrary family (or “path”) of values of (κ, t)* . This includes both sequences $((\kappa_n, t_n))_{n \in \mathbb{N}}$ and curves $((\kappa_s, t_s))_{s \in [0, \infty)}$, hence we omit subscripts. Without loss of generality, we assume that all the κ ’s have the same sign (just consider separately the subfamilies with positive and negative κ ’s). To simplify notation, we only consider positive families $\kappa \geq 0$ and give results for both κ and $-\kappa$.

Our main interest is for families of values of (κ, t) such that

$$\text{either } \kappa \rightarrow \infty \text{ with bounded } t, \quad \text{or } t \rightarrow 0 \text{ with arbitrary } \kappa \geq 0. \quad (2.3)$$

Whenever this holds, one has (see §A.1)

$$c(\kappa, t) \rightarrow 0, \quad p(-\kappa, t) \rightarrow 0. \quad (2.4)$$

We stress that (2.3) gathers many interesting regimes, namely:

- (a) $\kappa \rightarrow \infty$ and $t \rightarrow \bar{t} \in (0, \infty)$ (in particular, the case of fixed $t = \bar{t} > 0$);
- (b) $\kappa \rightarrow \infty$ and $t \rightarrow 0$;
- (c) $t \rightarrow 0$ and $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$ (in particular, the case of fixed $\kappa = \bar{\kappa} > 0$);
- (d) $t \rightarrow 0$ and $\kappa \rightarrow 0$.

Remarkably, while regime (d) needs to be handled separately, regimes (a)-(b)-(c) will be analyzed at once, as special instances of the case “ κ is bounded away from zero”.

Remark 2.1. We stress the requirement of *bounded* maturity t in (2.3). Some of our arguments can be adapted to deal with cases when $t \rightarrow \infty$, but additional work is needed (for instance, we assume the boundedness of some exponential moments $E[e^{(1+\eta)X_t}]$, cf. (2.9)-(2.10) below, which is satisfied by most models if t is bounded, but not if $t \rightarrow \infty$). We refer to [T09, JKM13] for results in the regime $t \rightarrow \infty$.

Given a model $(X_t)_{t \in [0, \infty)}$, the *implied volatility* $\sigma_{\text{imp}}(\kappa, t)$ is defined as the value of the volatility parameter $\sigma \in [0, \infty)$ that plugged into the Black&Scholes formula yields the given call and put prices $c(\kappa, t)$ and $p(\kappa, t)$ in (2.1) (see §4.2-§4.3 below). To avoid trivialities, we focus on families of (κ, t) such that $c(\kappa, t) > 0$ and $p(-\kappa, t) > 0$ (in fact, note that $\sigma_{\text{imp}}(\kappa, t) = 0$ if $c(\kappa, t) = 0$ and, likewise, $\sigma_{\text{imp}}(-\kappa, t) = 0$ if $p(-\kappa, t) = 0$).

Notation. Throughout the paper, we write $f(\kappa, t) \sim g(\kappa, t)$ to mean $f(\kappa, t)/g(\kappa, t) \rightarrow 1$. Let us recall a useful standard device (*subsequence argument*): to prove an asymptotic relation, such as e.g. $f(\kappa, t) \sim g(\kappa, t)$, it suffices to show that from every subsequence one can extract a further sub-subsequence along which the given relation holds. As a consequence, *in the proofs we may always assume that all quantities of interest have a (possibly infinite) limit*, e.g. $\kappa \rightarrow \bar{\kappa} \in [0, \infty]$ and $t \rightarrow \bar{t} \in [0, \infty)$, because this is true along a suitable subsequence.

2.2. Main results: atypical deviations. We first focus on families of (κ, t) such that

$$\bar{F}_t(\kappa) \rightarrow 0, \quad \text{resp.} \quad F_t(-\kappa) \rightarrow 0, \quad (2.5)$$

a regime that we call *atypical deviations*. This is the most interesting case, much studied in the literature, since it includes regimes (a), (b) and (c) described on page 2, and also regime (d) provided $\kappa \rightarrow 0$ sufficiently slow.

When $\kappa \rightarrow \infty$ with fixed $t > 0$, Benaim and Friz [BF09] require the *regular variation* of the tail probabilities, i.e. there exist $\alpha > 0$ and a slowly varying function[†] $L_t(\cdot)$ such that

$$\log \bar{F}_t(\kappa) \sim -L_t(\kappa) \kappa^\alpha, \quad \text{resp.} \quad \log F_t(-\kappa) \sim -L_t(\kappa) \kappa^\alpha. \quad (2.6)$$

It is not obvious how to generalize (2.6) when t is allowed to vary, i.e. which conditions to impose on $L_t(\kappa)$. However, one can reformulate the first relation in (2.6) simply requiring the existence of $\lim_{\kappa \rightarrow \infty} \log \bar{F}_t(\varrho\kappa) / \log \bar{F}_t(\kappa)$ for any fixed $\varrho > 0$, by [BGT89, Theorem 1.4.1], and analogously for the second relation in (2.6). This reformulation (in which $L_t(\kappa)$ is not even mentioned!) turns out to be the right condition to impose in the general context that we consider, when t is allowed to vary. We are thus led to the following:

Hypothesis 2.2 (Regular decay of tail probability). *The family of values of (κ, t) with $\kappa > 0$, $t > 0$ satisfies (2.5), and for every $\varrho \in [1, \infty)$ the following limit exists in $[0, +\infty]$:*

$$I_+(\varrho) := \lim \frac{\log \bar{F}_t(\varrho\kappa)}{\log \bar{F}_t(\kappa)}, \quad \text{resp.} \quad I_-(\varrho) := \lim \frac{\log F_t(-\varrho\kappa)}{\log F_t(-\kappa)}, \quad (2.7)$$

where limits are taken along the given family of values of (κ, t) . Moreover

$$\lim_{\varrho \downarrow 1} I_+(\varrho) = 1, \quad \text{resp.} \quad \lim_{\varrho \downarrow 1} I_-(\varrho) = 1. \quad (2.8)$$

Depending on the regime of κ , we will also need one of the following *moment conditions*.

- Given $\eta \in (0, \infty)$, the first moment condition is

$$\limsup E[e^{(1+\eta)X_t}] < \infty, \quad (2.9)$$

[†]A positive function $L(\cdot)$ is slowly varying if $\lim_{x \rightarrow \infty} L(\varrho x)/L(x) = 1$ for all $\varrho > 0$.

along the given family of values of (κ, t) . When $t \leq T$, it is enough to require that

$$\mathbb{E}[e^{(1+\eta)X_T}] < \infty, \quad (2.10)$$

because $(e^{(1+\eta)X_t})_{t \geq 0}$ is a submartingale and hence $\mathbb{E}[e^{(1+\eta)X_t}] \leq \mathbb{E}[e^{(1+\eta)X_T}]$.

- Given $\eta \in (0, \infty)$, the second moment condition is

$$\limsup \mathbb{E} \left[\left| \frac{e^{X_t} - 1}{\kappa} \right|^{1+\eta} \right] < \infty, \quad (2.11)$$

along the given family of values of (κ, t) . Note that for $\eta = 1$ this simplifies to

$$\exists C \in (0, \infty) : \quad \mathbb{E}[e^{2X_t}] \leq 1 + C\kappa^2. \quad (2.12)$$

We are ready to state our main results, which express the asymptotic behavior of option prices and implied volatility explicitly in terms of the tail probabilities. Due to different assumptions, we first consider right-tail asymptotics.

Theorem 2.3 (Right-tail atypical deviations). *Consider a family of values of (κ, t) with $\kappa > 0$, $t > 0$ such that Hypothesis 2.2 is satisfied by the right tail probability $\bar{F}_t(\kappa)$.*

- (i) **[κ bounded away from zero, t bounded away from infinity ($\liminf \kappa > 0$, $\limsup t < \infty$)]**
Let the moment condition (2.9) hold for every $\eta > 0$, or alternatively let it hold only for some $\eta > 0$ but in addition assume that

$$I_+(\varrho) \geq \varrho, \quad \forall \varrho \geq 1. \quad (2.13)$$

Then

$$\log c(\kappa, t) \sim \log \bar{F}_t(\kappa) + \kappa, \quad (2.14)$$

$$\sigma_{\text{imp}}(\kappa, t) \sim \left(\sqrt{\frac{-\log \bar{F}_t(\kappa)}{\kappa}} - \sqrt{\frac{-\log \bar{F}_t(\kappa)}{\kappa} - 1} \right) \sqrt{\frac{2\kappa}{t}}. \quad (2.15)$$

Special case: if $-\log \bar{F}_t(\kappa)/\kappa \rightarrow \infty$, assumption (2.13) can be relaxed to

$$\lim_{\varrho \rightarrow \infty} I_+(\varrho) = \infty, \quad (2.16)$$

and relations (2.14)-(2.15) simplify to

$$\log c(\kappa, t) \sim \log \bar{F}_t(\kappa), \quad (2.17)$$

$$\sigma_{\text{imp}}(\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log \bar{F}_t(\kappa))}}. \quad (2.18)$$

- (ii) **[κ and t vanish ($\kappa \rightarrow 0$, $t \rightarrow 0$)]** Let the moment condition (2.11) hold for every $\eta > 0$, or alternatively let it hold only for some $\eta > 0$ but in addition assume (2.16). Then

$$\log(c(\kappa, t)/\kappa) \sim \log \bar{F}_t(\kappa), \quad (2.19)$$

$$\sigma_{\text{imp}}(\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log \bar{F}_t(\kappa))}}. \quad (2.20)$$

Next we turn to left-tail asymptotics. The assumptions in this case turn out to be sensibly weaker than those for right-tail. For instance, the left-tail condition $\mathbb{E}[e^{-\eta X_T}] < \infty$ required in [BF09, Theorem 1.2] is not needed, which allows to treat the case of a *polynomially decaying* left tail, like in the Carr-Wu model described in Section 3.

Theorem 2.4 (Left-tail atypical deviations). *Consider a family of values of (κ, t) with $\kappa > 0, t > 0$ such that Hypothesis 2.2 is satisfied by the left tail probability $F_t(-\kappa)$.*

- $[\kappa$ bounded away from zero, t bounded away from infinity ($\liminf \kappa > 0, \limsup t < \infty$)]
With no moment condition and no extra assumption on $I_-(\cdot)$, one has

$$\log p(-\kappa, t) \sim \log F_t(-\kappa) - \kappa, \quad (2.21)$$

$$\sigma_{\text{imp}}(-\kappa, t) \sim \left(\sqrt{\frac{-\log F_t(-\kappa)}{\kappa}} + 1 - \sqrt{\frac{-\log F_t(-\kappa)}{\kappa}} \right) \sqrt{\frac{2\kappa}{t}}. \quad (2.22)$$

Special case: if $-\log F_t(-\kappa)/\kappa \rightarrow \infty$, relations (2.21)-(2.22) simplify to

$$\log p(-\kappa, t) \sim \log F_t(-\kappa), \quad (2.23)$$

$$\sigma_{\text{imp}}(-\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log F_t(-\kappa))}}. \quad (2.24)$$

- $[\kappa$ and t vanish ($\kappa \rightarrow 0, t \rightarrow 0$)] Let the moment condition (2.11) hold for every $\eta > 0$, or alternatively let it hold only for some $\eta > 0$ but in addition assume that

$$\lim_{\varrho \uparrow \infty} I_-(\varrho) = \infty. \quad (2.25)$$

Then

$$\log(p(-\kappa, t)/\kappa) \sim \log F_t(-\kappa), \quad (2.26)$$

$$\sigma_{\text{imp}}(-\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log F_t(-\kappa))}}. \quad (2.27)$$

We prove Theorems 2.3 and 2.4 in §5.1 below. The key step is to link the option prices $c(\kappa, t), p(\kappa, t)$ to the tail probabilities $\bar{F}_t(\kappa), F_t(-\kappa)$, exploiting Hypothesis 2.2. Once this is done, the asymptotic behavior of the implied volatility $\sigma_{\text{imp}}(\kappa, t)$ can be deduced using the model independent results of [GL14], that we summarize in §2.4.

Remark 2.5. The “special case” conditions

$$-\frac{\log \bar{F}_t(\kappa)}{\kappa} \rightarrow \infty, \quad \text{resp.} \quad -\frac{\log F_t(-\kappa)}{\kappa} \rightarrow \infty, \quad (2.28)$$

are automatically fulfilled in the small maturity regime $t \rightarrow 0$ with fixed strike $\kappa = \bar{\kappa} > 0$. In this case, one can use the simplified formulas (2.17)-(2.18) and (2.23)-(2.24).

2.3. Main results: typical deviations. Next we focus on the regime when $t \rightarrow 0$ and $\kappa \rightarrow 0$ sufficiently fast, so that the tail probability $\bar{F}_t(\kappa)$, resp. $F_t(-\kappa)$ has a strictly positive limit and condition (2.5) is violated. We call this regime *typical deviations*. This includes the basic regime of fixed $\kappa = 0$ and $t \downarrow 0$. Mixed regimes, when $\kappa \rightarrow 0$ and $t \rightarrow 0$ simultaneously, are also interesting, e.g. to interpolate between the at-the-money ($\kappa = 0$) and out-of-the-money ($\kappa \neq 0$) cases, which can be strikingly different as $t \rightarrow 0$ (see [MT12]).

We make the following natural assumption.

Hypothesis 2.6 (Small time scaling). *There is a positive function $(\gamma_t)_{t>0}$ with $\lim_{t \downarrow 0} \gamma_t = 0$ such that X_t/γ_t converges in law as $t \downarrow 0$ to some random variable Y :*

$$\frac{X_t}{\gamma_t} \xrightarrow[t \downarrow 0]{d} Y. \quad (2.29)$$

We refer to Remark 2.8 below for concrete ways to check Hypothesis 2.6. Let us stress that (2.29) is a condition on the tail probabilities, since it can be reformulated as

$$\bar{F}_t(a\gamma_t) \rightarrow \mathbb{P}(Y > a) \quad \text{and} \quad F_t(-a\gamma_t) \rightarrow \mathbb{P}(Y \leq -a), \quad (2.30)$$

for all $a \geq 0$ with $\mathbb{P}(|Y| = a) = 0$. If the support of the law of Y is unbounded from above and below (as it is usually the case), the limits in (2.30) are strictly positive for every $a \geq 0$.

The appropriate moment condition in this regime turns out to be (2.11) with $\kappa = \gamma_t$, i.e.

$$\exists \eta > 0 : \quad \limsup_{t \rightarrow 0} \mathbb{E} \left[\left| \frac{e^{X_t} - 1}{\gamma_t} \right|^{1+\eta} \right] < \infty. \quad (2.31)$$

Lastly, we introduce some notation. Denote by $\phi(\cdot)$ and $\Phi(\cdot)$ respectively the density and distribution function of a standard Gaussian (see (4.1) below), and define the function

$$D(z) := \frac{\phi(z)}{z} - \Phi(-z), \quad \forall z > 0. \quad (2.32)$$

As we show in §4.1 below, D is a smooth and strictly decreasing bijection from $(0, \infty)$ to $(0, \infty)$. Its inverse $D^{-1} : (0, \infty) \rightarrow (0, \infty)$ is also smooth, strictly decreasing and satisfies

$$D^{-1}(y) \sim \sqrt{2(-\log y)} \quad \text{as } y \downarrow 0, \quad D^{-1}(y) \sim \frac{1}{\sqrt{2\pi}} \frac{1}{y} \quad \text{as } y \uparrow \infty. \quad (2.33)$$

We can finally state the following result, linking option prices and implied volatility to tail probabilities in the regime of typical deviations.

Theorem 2.7 (Typical deviations). *Assume that Hypothesis 2.6 is satisfied, and moreover the moment condition (2.31) holds. Then the random variable in (2.29) satisfies $\mathbb{E}[Y] = 0$.*

Fix $a \in [0, \infty)$ with $\mathbb{P}(Y > a) > 0$, resp. $\mathbb{P}(Y < -a) > 0$. For any family of (κ, t) with

$$t \rightarrow 0 \quad \text{and} \quad \frac{\kappa}{\gamma_t} \rightarrow a \in [0, \infty),$$

the asymptotic behavior of option prices is given by

$$c(\kappa, t) \sim \gamma_t \mathbb{E}[(Y - a)^+], \quad \text{resp.} \quad p(-\kappa, t) \sim \gamma_t \mathbb{E}[(Y + a)^-], \quad (2.34)$$

and correspondingly the implied volatility is given by

$$\sigma_{\text{imp}}(\pm\kappa, t) \sim C_{\pm}(a) \frac{\gamma_t}{\sqrt{t}}, \quad \text{with} \quad C_{\pm}(a) = \begin{cases} \frac{a}{D^{-1}(\frac{\mathbb{E}[(Y \mp a)^{\pm}]}{a})} & \text{if } a > 0, \\ \sqrt{2\pi} \mathbb{E}[Y^{\pm}] & \text{if } a = 0. \end{cases} \quad (2.35)$$

Remark 2.8. Hypothesis 2.6 can be easily checked when the characteristic function of X_t is known, because, by Lévy continuity theorem, the convergence in distribution (2.29) is equivalent to the pointwise convergence $\mathbb{E}[e^{iuX_t/\gamma_t}] \rightarrow \mathbb{E}[e^{iuY}]$ for every $u \in \mathbb{R}$. We will see concrete examples in Subsections 3.1 (Carr-Wu model) and 3.2 (Merton's model).

Another interesting case is that of *diffusions*. Assume that $X_t = \log S_t$, where $(S_t)_{t \geq 0}$ evolves according to the stochastic differential equation

$$\begin{cases} dS_t = \sqrt{V_t} S_t dW_t \\ S_0 = 1 \end{cases}, \quad (2.36)$$

where $W = (W_t)_{t \geq 0}$ is a Brownian motion and $V = (V_t)_{t \geq 0}$ is a positive adapted process, representing the volatility, possibly correlated with W . Under the mild assumption that $\lim_{t \rightarrow 0} V_t = \sigma_0^2$ a.s., where $\sigma_0 \in (0, \infty)$ is a *constant*, one can show that Hypothesis 2.6 holds with $\gamma_t = \sqrt{t}$ and $Y \sim N(0, \sigma_0^2)$, see Appendix A.2.

Interestingly, plugging $Y \sim N(0, \sigma_0^2)$ into (2.35) yields $C_{\pm}(a) \equiv \sigma_0$ (see Appendix A.2). Consequently, if the moment condition (2.31) holds, we can apply Theorem 2.7, getting $\sigma_{\text{imp}}(\pm\kappa, t) \sim \sigma_0$ along any parabolic curve $\kappa \sim a\sqrt{t}$. This result is consistent with recent results by Pagliarani and Pascucci [PP15], who go beyond first-order asymptotics.

The proof of Theorem 2.7 is given in §5.2 below. The asymptotic behavior (2.34) of option prices follows easily from the convergence in distribution (2.29), because the needed uniform integrability is ensured by the moment condition (2.31). The asymptotic behavior (2.35) of implied volatility can again be deduced from the option prices asymptotics in a model independent way, that we now describe.

2.4. From option price to implied volatility. Whenever the option prices $c(\kappa, t)$ or $p(-\kappa, t)$ vanish, they determine the asymptotic behavior of the implied volatility through *explicit universal formulas*. These are summarized in the following theorem (of which we give in Section 4 a self-contained proof), which gathers results from the recent literature.

Theorem 2.9 (From option price to implied volatility). *Consider an arbitrary family of values of (κ, t) with $\kappa \geq 0$ and $t > 0$, such that $c(\kappa, t) \rightarrow 0$, resp. $p(-\kappa, t) \rightarrow 0$.*

- Case of κ bounded away from zero (i.e. $\liminf \kappa > 0$).

$$\begin{aligned}\sigma_{\text{imp}}(\kappa, t) &\sim \left(\sqrt{\frac{-\log c(\kappa, t)}{\kappa}} + 1 - \sqrt{\frac{-\log c(\kappa, t)}{\kappa}} \right) \sqrt{\frac{2\kappa}{t}}, \quad \text{resp.} \\ \sigma_{\text{imp}}(-\kappa, t) &\sim \left(\sqrt{\frac{-\log p(-\kappa, t)}{\kappa}} - \sqrt{\frac{-\log p(-\kappa, t)}{\kappa} - 1} \right) \sqrt{\frac{2\kappa}{t}}.\end{aligned}\tag{2.37}$$

- Case of $\kappa \rightarrow 0$, with $\kappa > 0$.

$$\begin{aligned}\sigma_{\text{imp}}(\kappa, t) &\sim \frac{1}{D^{-1}\left(\frac{c(\kappa, t)}{\kappa}\right)} \frac{\kappa}{\sqrt{t}}, \quad \text{resp.} \\ \sigma_{\text{imp}}(-\kappa, t) &\sim \frac{1}{D^{-1}\left(\frac{p(-\kappa, t)}{\kappa}\right)} \frac{\kappa}{\sqrt{t}},\end{aligned}\tag{2.38}$$

where the function $D : (0, \infty) \rightarrow (0, \infty)$ is defined in (2.32)-(2.33).

- Case of $\kappa = 0$.

$$\sigma_{\text{imp}}(0, t) \sim \sqrt{2\pi} \frac{c(0, t)}{\sqrt{t}} = \sqrt{2\pi} \frac{p(0, t)}{\sqrt{t}}.\tag{2.39}$$

We stress that Theorem 2.9 allows to derive immediately all the asymptotic relations for the implied volatility $\sigma_{\text{imp}}(\pm\kappa, t)$ appearing in Theorems 2.3, 2.4 and 2.7 from the corresponding relations for the option prices $c(\kappa, t)$ and $p(-\kappa, t)$.

The main part of Theorem 2.9 is equation (2.37), which was recently proved by Gao and Lee [GL14] extending previous results of Lee [L04], Roper and Rutkowski [RR09], Benaim and Friz [BF09] and Gulisashvili [G10]. As a matter of fact, Gao and Lee prove much more than (2.37), providing explicit estimates for the error beyond first order asymptotics.

Equation (2.38) is a new contribution of the present paper. In fact, [GL14] assume that $-\log \kappa = o(-\log c(\kappa, t))$ (cf. equation (4.2) therein), which excludes the regimes with $\kappa \rightarrow 0$ “fast enough”. The relevance of such regimes has been recently shown in [MT12], where the special case $\kappa \propto \sqrt{t \log(1/t)}$ is considered (see [MT12, Theorem 3.1]).

Remark 2.10. Relation (2.38) provides an *interpolation between the at-the-money and out-of-the-money regimes*, described by (2.39) and (2.37). Let us be more explicit.

Using (2.33), formula (2.38) can be rewritten as follows:

$$\sigma_{\text{imp}}(\kappa, t) \sim \begin{cases} \frac{\kappa}{\sqrt{2t(-\log(c(\kappa, t)/\kappa))}} & \text{if } \frac{c(\kappa, t)}{\kappa} \rightarrow 0; \\ \frac{\kappa}{D^{-1}(a)\sqrt{t}} & \text{if } \frac{c(\kappa, t)}{\kappa} \rightarrow a \in (0, \infty); \\ \sqrt{2\pi} \frac{c(\kappa, t)}{\sqrt{t}} & \text{if } \frac{c(\kappa, t)}{\kappa} \rightarrow \infty, \text{ or if } \kappa = 0, \end{cases} \quad (2.40)$$

and analogously for $\sigma_{\text{imp}}(-\kappa, t)$, just replacing $c(\kappa, t)$ by $p(-\kappa, t)$.

Note that the last line in (2.40) matches with the at-the-money regime (2.39). In order to see how (2.40) matches with the out-of-the-money regime (2.37), it suffices to note that whenever $\frac{-\log c(\kappa, t)}{\kappa} \rightarrow \infty$, resp. $\frac{-\log p(-\kappa, t)}{\kappa} \rightarrow \infty$, formula (2.37) can be rewritten as

$$\sigma_{\text{imp}}(\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log c(\kappa, t))}}, \quad \text{resp.} \quad \sigma_{\text{imp}}(-\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log p(-\kappa, t))}}, \quad (2.41)$$

and this coincides with the first line of (2.40) when $\kappa \rightarrow 0$ slowly enough, namely

$$-\log \kappa = o(-\log c(\kappa, t)). \quad (2.42)$$

2.5. Discussion. Theorems 2.3, 2.4 and 2.7 are useful because their assumptions, involving asymptotics for the tail probabilities $\bar{F}_t(\kappa)$ and $F_t(-\kappa)$, can be directly verified for many concrete models (see Section 3 for some examples). The difference between the regimes of typical and atypical deviations can be described as follows:

- for typical deviations, the key assumption is Hypothesis 2.6, which concerns the *weak convergence* of X_t , cf. (2.29)-(2.30);
- for atypical deviations, the key assumption is Hypothesis 2.2, which concerns the *large deviations* properties of X_t , cf. (2.7)-(2.8).

In particular, it is worth stressing that Hypothesis 2.2 requires sharp asymptotics only for the *logarithm of the tail probabilities* $\log \bar{F}_t(\kappa)$ and $\log F_t(-\kappa)$, and not for the tail probabilities themselves, which would be a considerably harder task (out of reach for many models). As a consequence, Hypothesis 2.2 can often be checked through the celebrated *Gärtner-Ellis Theorem* [DZ98, Theorem 2.3.6], which yields sharp asymptotics on $\log \bar{F}_t(\kappa)$ and $\log F_t(-\kappa)$ under suitable conditions on the moment generating function of X_t .

3. APPLICATIONS

In this section we show the relevance of our main theoretical results, deriving asymptotic expansions of the implied volatility for Carr-Wu finite moment logstable model (§3.1) and Merton's jump diffusion model (§3.2). The case of Heston's model is briefly discussed in §3.3.

Our results can also be applied to a stochastic volatility model, recently introduced in [ACDP12], which exhibits multiscaling of moments. Even though no closed expression is available for the moment generating function of the log-price, the tail probabilities can be estimated *explicitly*, as we show in a separate paper [CC15]. This leads to precise asymptotics for the implied volatility, thanks to Theorems 2.3, 2.4 and 2.7.

3.1. Carr-Wu Finite Moment Logstable Model. Carr and Wu [CW04] consider a model where the log-strike X_t has characteristic function

$$\mathbb{E}[e^{iuX_t}] = e^{t[iu\mu - |u|^\alpha \sigma^\alpha (1 + i \operatorname{sign}(u) \tan(\frac{\pi\alpha}{2}))]}, \quad (3.1)$$

where $\sigma \in (0, \infty)$, $\alpha \in (1, 2]$ and we fix $\mu := \sigma^\alpha / \cos(\frac{\pi\alpha}{2})$ to work in the risk-neutral measure, cf. [CW04, Proposition 1]. The moment generating function of X_t is

$$\mathbb{E}[e^{\lambda X_t}] = \begin{cases} e^{[\lambda\mu - \frac{(\lambda\sigma)^\alpha}{\cos(\frac{\pi\alpha}{2})}]t} & \text{if } \lambda \geq 0, \\ +\infty & \text{if } \lambda < 0. \end{cases} \quad (3.2)$$

Note that as $\alpha \rightarrow 2$ one recovers Black&Scholes model with volatility $\sqrt{2}\sigma$, cf. §4.2 below.

Applying Theorems 2.3, 2.4 and 2.7, we give a *complete characterization* of the volatility smile asymptotics with bounded maturity. This includes, in particular, the regimes of extreme strike ($\kappa \rightarrow \pm\infty$ with fixed $t > 0$) and of small maturity ($t \rightarrow 0$ with fixed κ).

Theorem 3.1 (Smile asymptotics of Carr-Wu model). *The following asymptotics hold.*

- **Atypical deviations.** Consider any family of (κ, t) with $\kappa \geq 0$, $t > 0$ such that

$$\text{either } t \rightarrow 0 \text{ and } \kappa \gg t^{1/\alpha}, \quad \text{or } t \rightarrow \bar{t} \in (0, \infty) \text{ and } \kappa \rightarrow \infty. \quad (3.3)$$

(This includes the regimes (a), (b), (c) on page 2, and part of regime (d).) Then one has the right-tail asymptotics

$$\sigma_{\text{imp}}(\kappa, t) \sim B_\alpha \left(\frac{\kappa}{t} \right)^{-\frac{2-\alpha}{2(\alpha-1)}}, \quad \text{where } B_\alpha := \frac{(\alpha\sigma)^{\alpha/2}}{\sqrt{2(\alpha-1)} |\cos(\frac{\pi\alpha}{2})|^{\frac{1/2}{\alpha-1}}}. \quad (3.4)$$

The corresponding left-tail asymptotics are given by

$$\sigma_{\text{imp}}(-\kappa, t) \sim \left(\sqrt{\frac{\log \frac{\kappa^\alpha}{t}}{\kappa}} + 1 - \sqrt{\frac{\log \frac{\kappa^\alpha}{t}}{\kappa}} \right) \sqrt{\frac{2\kappa}{t}}, \quad (3.5)$$

which can be made more explicit distinguishing different regimes:

$$\sigma_{\text{imp}}(-\kappa, t) \sim \begin{cases} \frac{\kappa}{\sqrt{2t \log \frac{\kappa^\alpha}{t}}} & \text{if } t \rightarrow 0 \text{ and } \frac{\kappa}{\log \frac{1}{t}} \rightarrow 0, \\ \left(\frac{\sqrt{1+a}-1}{\sqrt{a}} \right) \sqrt{\frac{2\kappa}{t}} & \text{if } t \rightarrow 0 \text{ and } \frac{\kappa}{\log \frac{1}{t}} \rightarrow a \in (0, \infty), \\ \sqrt{\frac{2\kappa}{t}} & \begin{cases} \text{if } t \rightarrow 0 \text{ and } \frac{\kappa}{\log \frac{1}{t}} \rightarrow \infty, \\ \text{if } t \rightarrow \bar{t} \in (0, \infty) \text{ and } \kappa \rightarrow \infty. \end{cases} \end{cases}. \quad (3.6)$$

- **Typical deviations.** For any family of (κ, t) with

$$t \rightarrow 0, \quad \frac{\kappa}{t^{1/\alpha}} \rightarrow a \in [0, \infty), \quad (3.7)$$

one has

$$\sigma_{\text{imp}}(\pm\kappa, t) \sim C_\pm(a) t^{\frac{2-\alpha}{2\alpha}}, \quad \text{with } C_\pm(a) := \begin{cases} \frac{a}{D^{-1}(\frac{\mathbb{E}[(\sigma Y \mp a)^\pm]}{a})} & \text{if } a > 0, \\ \sqrt{2\pi} \sigma \mathbb{E}[Y^\pm] & \text{if } a = 0. \end{cases} \quad (3.8)$$

Remark 3.2 (Surface asymptotics for the Carr-Wu model). The fact that relations (3.4) and (3.5) hold for *any* family of (κ, t) satisfying (3.3) yields interesting consequences. We claim that, for any $T \in (0, \infty)$ and $\varepsilon > 0$, there exists $M = M(\varepsilon, T) \in (0, \infty)$ such that the following inequalities hold *for all* (κ, t) in the region $\mathcal{A}_{T,M} := \{0 < t \leq T, \kappa > Mt^{1/\alpha}\}$:

$$(1 - \varepsilon) B_\alpha \left(\frac{\kappa}{t} \right)^{-\frac{2-\alpha}{2(\alpha-1)}} \leq \sigma_{\text{imp}}(\kappa, t) \leq (1 + \varepsilon) B_\alpha \left(\frac{\kappa}{t} \right)^{-\frac{2-\alpha}{2(\alpha-1)}}. \quad (3.9)$$

Similar inequalities can be deduced from (3.5)-(3.6) and (3.8). Relation (3.9) gives a *uniform approximations* of the volatility surface $\sigma_{\text{imp}}(\kappa, t)$ in open regions of the plane (κ, t) .

The proof of (3.9) is simple: assume by contradiction that there exist $T, \varepsilon \in (0, \infty)$ such that for every $M \in (0, \infty)$ relation (3.9) *fails* for some $(\kappa_M, t_M) \in \mathcal{A}_{T,M}$. Extracting a subsequence, the family (κ_M, t_M) satisfies (3.3) but not (3.4), contradicting Theorem 3.1.

Proof of Theorem 3.1. Let Y denote a random variable with characteristic function

$$\mathbb{E}[e^{iuY}] = e^{-|u|^\alpha (1 + i \text{sign}(u) \tan(\frac{\pi\alpha}{2}))}, \quad (3.10)$$

i.e. Y has a strictly stable law with index α and skewness parameter $\beta = -1$, and $\mathbb{E}[Y] = 0$.

If we set

$$Y_t := \frac{X_t - \mu t}{\sigma t^{1/\alpha}}, \quad (3.11)$$

it follows by (3.1) that Y_t has the same distribution as Y , because

$$\mathbb{E}[e^{iuY_t}] = \mathbb{E}[e^{iuY}] = e^{-|u|^\alpha (1 + i \text{sign}(u) \tan(\frac{\pi\alpha}{2}))}. \quad (3.12)$$

It follows by (3.11) that

$$\frac{X_t}{t^{1/\alpha}} \xrightarrow[t \downarrow 0]{d} \sigma Y, \quad (3.13)$$

hence Hypothesis 2.6 is satisfied with $\gamma_t := t^{1/\alpha}$.

Note that $\mathbb{P}(Y > a) > 0$ and $\mathbb{P}(Y < -a) > 0$ for all $a \in \mathbb{R}$, because the density of Y is strictly positive everywhere. The right tail of Y has a super-exponential decay: as $\kappa \rightarrow \infty$

$$\log \mathbb{P}(Y > \kappa) \sim -\tilde{B}_\alpha \kappa^{\alpha/(\alpha-1)} \quad \text{where} \quad \tilde{B}_\alpha := \frac{\alpha-1}{\alpha} \left(\frac{|\cos(\frac{\pi\alpha}{2})|}{\alpha} \right)^{1/(\alpha-1)}, \quad (3.14)$$

cf. [CW04, Property 1 and references therein]. On the other hand the left tail is polynomial: there exists $c = c_\alpha \in (0, \infty)$ such that

$$\mathbb{P}(Y \leq -\kappa) \sim \frac{c}{\kappa^\alpha}, \quad \text{hence} \quad \log \mathbb{P}(Y \leq -\kappa) \sim -\alpha \log \kappa. \quad (3.15)$$

Recalling that $\bar{F}_t(\kappa) := \mathbb{P}(X_t > \kappa)$ and $F_t(-\kappa) := \mathbb{P}(X_t \leq \kappa)$, by (3.11) we can write

$$\bar{F}_t(\kappa) = \mathbb{P}\left(Y > \frac{\kappa - \mu t}{\sigma t^{1/\alpha}}\right), \quad F_t(-\kappa) = \mathbb{P}\left(Y \leq \frac{-\kappa - \mu t}{\sigma t^{1/\alpha}}\right), \quad (3.16)$$

hence we can transfer the estimates (3.14) and (3.15) to X_t .

Henceforth we consider separately the regimes of atypical deviations (3.3), and that of typical deviations (3.7). Note that it is easy to check that (3.5) is equivalent to (3.6).

Atypical deviations. Let us fix an arbitrary family of values of (κ, t) satisfying (3.3). Then $\kappa/t \rightarrow \infty$ (because $\alpha > 1$), hence

$$\frac{\kappa - \mu t}{\sigma t^{1/\alpha}} \sim \frac{\kappa}{\sigma t^{1/\alpha}} \rightarrow \infty, \quad \frac{-\kappa - \mu t}{\sigma t^{1/\alpha}} \sim \frac{-\kappa}{\sigma t^{1/\alpha}} \rightarrow -\infty.$$

By (3.14), (3.15) and (3.16) we then obtain

$$\log \bar{F}_t(\kappa) \sim -\tilde{B}_\alpha \left(\frac{\kappa}{\sigma t^{1/\alpha}} \right)^{\alpha/(\alpha-1)}, \quad \log F_t(-\kappa) \sim -\log \frac{\kappa^\alpha}{t}. \quad (3.17)$$

Let us now check the assumptions of Theorem 2.3.

- The first relation in (3.17) shows that Hypothesis 2.2 is satisfied by the right tail $\bar{F}_t(\kappa)$, with $I_+(\varrho) = \varrho^{\alpha/(\alpha-1)}$. Note that $I_+(\varrho) \geq \varrho$ for all $\varrho \geq 1$, since $\alpha > 1$, hence also condition (2.13) is satisfied.
- Condition (2.9) is satisfied because (2.10) holds for all $T > 0$ and $\eta > 0$, by (3.2).
- It remains to check condition (2.11). As we prove below, for all $\eta \in (0, \alpha - 1)$ and $T > 0$ there are constants $A, B, C \in (0, \infty)$, depending on η, T and on the parameters α, σ , such that for all $0 < t \leq T$ and $\kappa \geq 0$ the following inequality holds:

$$\mathbb{E} \left[\left| \frac{e^{X_t} - 1}{\kappa} \right|^{1+\eta} \right] \leq A \left(\left(\frac{t^{1/\alpha}}{\kappa} \right)^B + C \right). \quad (3.18)$$

In particular, since $\kappa/t^{1/\alpha} \rightarrow \infty$ by assumption (3.3), condition (2.11) is satisfied.

Applying Theorem 2.3, since $-\log \bar{F}_t(\kappa)/\kappa \rightarrow \infty$ by the first relation in (3.17), the asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ is given by (2.18), which by (3.17) coincides with (3.4).

Next we want to apply Theorem 2.4. By the second relation in (3.17), Hypothesis 2.2 is satisfied by the left tail $F_t(-\kappa)$, with $I_-(\varrho) \equiv 1$. If κ is bounded away from zero, the asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ is given by (2.22), which by (3.17) yields precisely (3.5).

If $\kappa \rightarrow 0$ we cannot apply directly Theorem 2.4, because the moment condition (2.11) is satisfied only for some $\eta > 0$, and condition (2.25) is not satisfied, since $I_-(\varrho) \equiv 1$. However, we can show that (2.26) still holds by direct estimates. By (2.1)

$$p(-\kappa, t) = \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t < -\kappa\}}] \geq \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t < -2\kappa\}}] \geq (e^{-\kappa} - e^{-2\kappa}) F_t(-2\kappa),$$

and since $(e^{-\kappa} - e^{-2\kappa}) = e^{-2\kappa}(e^\kappa - 1) \geq e^{-2\kappa}\kappa$, we can write by (3.17) (recall that $\kappa \rightarrow 0$)

$$\log(p(-\kappa, t)/\kappa) \geq -2\kappa - \log \frac{(2\kappa)^\alpha}{t} \sim -\log \frac{\kappa^\alpha}{t}. \quad (3.19)$$

Next we give a matching upper bound on $p(-\kappa, t)$. Since $\mu t \leq \kappa$ eventually (recall that $\kappa/t^{1/\alpha} \rightarrow \infty$, hence $\kappa/t \rightarrow \infty$), by (3.16) and (3.15) we obtain, for all $y \geq 1$

$$F_t(-\kappa y) \leq \mathbb{P} \left(Y \leq -\frac{2\kappa y}{\sigma t^{1/\alpha}} \right) \leq c' \frac{t}{\kappa^\alpha y^\alpha},$$

for some $c' = c'_{\alpha, \sigma, \mu} \in (0, \infty)$. Then by Fubini's theorem

$$\begin{aligned} p(-\kappa, t) &= \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t < -\kappa\}}] = \mathbb{E} \left[\int_{\kappa}^{\infty} e^{-x} \mathbf{1}_{\{x < -X_t\}} dx \right] = \int_{\kappa}^{\infty} e^{-x} F_t(-x) dx \\ &= \kappa \int_1^{\infty} e^{-\kappa y} F_t(-\kappa y) dy \leq c' \kappa \frac{t}{\kappa^\alpha} \int_1^{\infty} \frac{1}{y^\alpha} dy =: c'' \kappa \frac{t}{\kappa^\alpha}, \end{aligned}$$

hence

$$\log(p(-\kappa, t)/\kappa) \leq \log c'' - \log \frac{\kappa^\alpha}{t} \sim -\log \frac{\kappa^\alpha}{t}.$$

This relation, together with (3.19), yields

$$\log(p(-\kappa, t)/\kappa) \sim -\log \frac{\kappa^\alpha}{t}.$$

Since $\kappa/t^{1/\alpha} \rightarrow \infty$, this shows that we are in the regime when $\kappa \rightarrow 0$ and $p(-\kappa, t)/\kappa \rightarrow 0$. We can thus apply equation (2.38) in Theorem 2.9, which recalling Remark 2.10 simplifies as the first line in (2.40) (with $p(-\kappa, t)$ instead of $c(\kappa, t)$), yielding

$$\sigma_{\text{imp}}(-\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log(p(-\kappa, t)/\kappa))}} \sim \frac{\kappa}{\sqrt{2t \log \frac{\kappa^\alpha}{t}}},$$

hence (3.5) is proved also when $\kappa \rightarrow 0$ (cf. (3.6)).

Typical deviations. Let us fix an arbitrary family of values of (κ, t) satisfying (3.7). Relation (3.18) for $\kappa = \gamma_t = t^{1/\alpha}$ shows that condition (2.31) is satisfied, and Hypothesis 2.6 holds by (3.13). We can then apply Theorem 2.7, and relation (2.35) gives precisely (3.8). \square

Proof of (3.18). Since $|\frac{e^x-1}{x}| \leq 1$ if $x < 0$ and $|\frac{e^x-1}{x}| \leq e^x$ if $x \geq 0$, we have $|\frac{e^x-1}{x}| \leq 1 + e^x$ for all $x \in \mathbb{R}$. If $p, q > 1$ are such that $\frac{1}{p} + \frac{1}{q} = 1$, Young's inequality $ab \leq \frac{1}{p}a^p + \frac{1}{q}b^q$ yields

$$\left| \frac{e^{X_t} - 1}{\kappa} \right| = \left| \frac{X_t}{\kappa} \right| \left| \frac{e^{X_t} - 1}{X_t} \right| \leq \frac{1}{p} \left(\frac{|X_t|}{\kappa} \right)^p + \frac{1}{q} (1 + e^{X_t})^q.$$

Noting that $(a + b)^r \leq 2^{r-1}(a^r + b^r)$ for $r \geq 1$, by Hölder's inequality, and denoting by $c = c_{p,\eta}$ a suitable constant depending only on p, η , we can write

$$\left| \frac{e^{X_t} - 1}{\kappa} \right|^{1+\eta} \leq c \left(\frac{|X_t|^{p(1+\eta)}}{\kappa^{p(1+\eta)}} + 1 + e^{q(1+\eta)X_t} \right).$$

Given $0 < \eta < \alpha - 1$, we fix $p = p_{\eta,\alpha} > 1$ such that $B := p(1 + \eta) < \alpha$. (Note that B depends only on η, α .) Moreover, it follows by (3.11) that

$$\mathbb{E}[|X_t|^B] = (\sigma t^{1/\alpha})^B \mathbb{E}[|Y|^B] (1 + \mathcal{O}(t^{B(1-1/\alpha)})),$$

and note that $\mathbb{E}[|Y|^B] < \infty$, because Y has finite moments of all orders strictly less than α , cf. [CW04, Property 1]. Since for $t \leq T$ one has $\mathbb{E}[e^{q(1+\eta)X_t}] \leq \mathbb{E}[e^{q(1+\eta)X_T}] < \infty$, by (3.2), relation (3.18) is proved. \square

3.2. Merton's Jump Diffusion Model. Consider a model [M76] where the log-return X_t has an infinitely divisible distribution, whose moment generating function is given by

$$\mathbb{E}[\exp(zX_t)] = \exp \left(t \left\{ z\mu + \frac{1}{2}z^2\sigma^2 + \lambda \left(e^{z\alpha+z^2\frac{\delta^2}{2}} - 1 \right) \right\} \right), \quad \forall z \in \mathbb{C}, \quad (3.20)$$

where $\mu, \alpha \in \mathbb{R}$ and $\sigma, \lambda, \delta \in (0, \infty)$ are fixed parameters.

The asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ has been studied by many authors. The case of fixed $t > 0$ and $\kappa \rightarrow \infty$ was derived by Benaim and Friz [BF09] using saddle point methods (for the detailed computation see [FGY14], [GMZ14]). The case of fixed $\kappa > 0$ and $t \rightarrow 0$ follows by [FF12], while the mixed regime of $t \rightarrow 0$, $\kappa \rightarrow 0$ with $\kappa \propto \sqrt{t \log(1/t)}$ was considered in [MT12]. Applying our results, we can complete the picture, providing general formulas which interpolate between all these regimes, cf. Theorem 3.4.

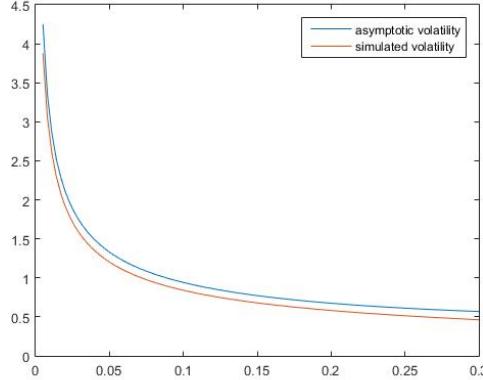


FIGURE 1. Implied volatility $\sigma_{\text{imp}}(\kappa, t)$ for Merton's model (with parameters $\lambda = 0.01$, $\delta = 0.3$, $\sigma = 0.2$, $\alpha = 0.1$) in the regime $\kappa = \kappa_2(t)$ for $t \in (0, 0.3)$. The *asymptotic volatility* is given by our formula (3.26), while the *simulated volatility* is obtained using standard computational packages.

Let us define two functions $\kappa_1(t) \rightarrow 0$ and $\kappa_2(t) \rightarrow \infty$ as $t \rightarrow 0$ as follows:

$$\kappa_1(t) := \sqrt{t \log \frac{1}{t}}, \quad \kappa_2(t) := \sqrt{\log \frac{1}{t}}, \quad (3.21)$$

which will separate different behaviors as $t \rightarrow 0$. We stress that $\kappa_1(t)$ is precisely the scaling considered in [MT12]. Let us also define $f : [0, \infty) \rightarrow (0, \infty)$ by

$$f(a) := \min_{n \in \mathbb{N}} \left(n + \frac{a^2}{2n\delta^2} \right). \quad (3.22)$$

Note that f is continuous and piecewise quadratic: more precisely, by explicit computation, $f(a) = n + \frac{a^2}{2n\delta^2}$ for all $a \in [\sqrt{2(n-1)n}\delta, \sqrt{2n(n+1)}\delta)$, with $n \in \mathbb{N}$. It follows that

$$f(0) = 1, \quad f(a) \underset{a \rightarrow \infty}{\sim} \frac{\sqrt{2}}{\delta} a. \quad (3.23)$$

The role of the function f is explained by the following Lemma, proved in Appendix A.3.

Lemma 3.3. *For every fixed $a \in (0, \infty)$, as $t \rightarrow 0$ one has*

$$\log P(X_t > a \kappa_2(t)) \sim -f(a) \log \frac{1}{t}. \quad (3.24)$$

Moreover, if either $t \rightarrow 0$ and $\kappa \gg \kappa_2(t)$, or if $t \rightarrow \bar{t} \in (0, \infty)$ and $\kappa \rightarrow \infty$, one has

$$\log P(X_t > \kappa) \sim -\frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}}. \quad (3.25)$$

We are now ready to state our main result for Merton's model (see Figure 1).

Theorem 3.4 (Smile Asymptotics of Merton's model). *Consider a family of values of (κ, t) with $\kappa \geq 0$ and $t > 0$.*

(1) *If $t \rightarrow 0$ and $\kappa = \mathcal{O}(\kappa_2(t))$, then*

$$\sigma_{\text{imp}}(\kappa, t) \sim \max \left\{ \sigma, \frac{\kappa}{\sqrt{2t f\left(\frac{\kappa}{\kappa_2(t)}\right) \log \frac{\kappa}{t}}} \right\}, \quad (3.26)$$

which can be rewritten more explicitly as follows:

$$\sigma_{\text{imp}}(\kappa, t) \sim \begin{cases} \sigma & \text{if } 0 \leq \kappa \leq \sigma \kappa_1(t) \\ \frac{\kappa}{\sqrt{2t \log \frac{\kappa}{t}}} & \text{if } \sigma \kappa_1(t) \leq \kappa \ll \kappa_2(t) \\ \frac{\kappa}{\sqrt{2t f(a) \log \frac{\kappa}{t}}} & \text{if } \kappa \sim a \kappa_2(t) \text{ with } a \in (0, \infty) \end{cases}. \quad (3.27)$$

(2) If $t \rightarrow 0$ and $\kappa \gg \kappa_2(t)$, or if $t \rightarrow \bar{t} \in (0, \infty)$ and $\kappa \rightarrow \infty$, then

$$\sigma_{\text{imp}}(\kappa, t) \sim \sqrt{\frac{\delta \kappa}{2t \sqrt{2 \log \frac{\kappa}{t}}}}. \quad (3.28)$$

Proof. We have to prove relation (3.27) (which is equivalent to (3.26), by extracting subsequences) and relation (3.28). We distinguish different subcases.

Assume first that $t \rightarrow 0$ with $\kappa = \mathcal{O}(\sqrt{t})$. By extracting sub-sequences, assume that $\kappa/\sqrt{t} \rightarrow a \in [0, \infty)$. Note that $X_t/\sqrt{t} \xrightarrow{d} N(0, \sigma^2)$, because $\mathbb{E}[e^{iu \frac{X_t}{\sqrt{t}}}] \rightarrow e^{-\frac{u^2}{2\sigma^2}}$ for every $u \in \mathbb{R}$, as one checks by (3.20). We then apply Theorem 2.7 with $\gamma_t = \sqrt{t}$, because the moment condition (2.31) for $\eta = 1$ follows by (3.20) (see also (2.12)). Relation (2.35) yields

$$\sigma(\kappa, t) \sim C_+(a) \frac{\sqrt{t}}{\sqrt{t}} = \sigma,$$

since $C_+(a) \equiv \sigma$ (see (A.4) in Appendix A.2). This matches with the first line of (3.27).

Next we assume that $t \rightarrow 0$ with $\sqrt{t} \ll \kappa \ll 1$. Applying [MT12, Proposition 2.3], we can write

$$c(\kappa, t) \sim \mathbb{E}[(e^{\sigma W_t - \frac{\sigma^2 t}{2}} - e^\kappa)^+] + C t,$$

with $0 < C := \int_0^\infty (e^x - 1) \nu(dx) < \infty$, where ν denotes the Lévy measure of X . The first term is the usual Black&Scholes price of a call option with $\kappa \gg \sqrt{t}$: applying (4.12) below with $v = \sigma \sqrt{t}$ and $d_1 = -\frac{\kappa}{v} + \frac{v}{2} \sim -\frac{\kappa}{\sigma \sqrt{t}}$, together with (4.1) and (4.3), we get

$$\frac{c(\kappa, t)}{\kappa} \sim \frac{v}{\kappa} \frac{\phi(d_1)}{(d_1)^2} + C \frac{t}{\kappa} = \frac{e^{-\frac{d_1^2}{2}}}{\sqrt{2\pi} (\frac{\kappa^2}{\sigma^2 t})^{3/2}} + C \frac{t}{\kappa} \sim \frac{e^{-\frac{d_1^2}{2}}}{\sqrt{2\pi} d_1^3} + C \frac{t}{\kappa}.$$

Writing $\frac{e^{-\frac{z^2}{2}}}{\sqrt{2\pi} z^3} = e^{-\frac{z^2}{2} - \log(\sqrt{2\pi} z^3)} = e^{-\frac{z^2}{2}(1+o(1))}$ as $z \rightarrow \infty$, we get

$$\frac{c(\kappa, t)}{\kappa} \sim e^{-\frac{\kappa^2}{2\sigma^2 t}(1+o(1))} + C \frac{t}{\kappa} = a + b \quad (\text{say}).$$

The inequalities $\max\{a, b\} \leq a + b \leq 2 \max\{a, b\}$ yield $\log(a + b) \sim \max\{\log a, \log b\}$ (the additive constant $\log 2$ is irrelevant, since $a, b \rightarrow 0$) hence

$$-\log \frac{c(\kappa, t)}{\kappa} \sim -\max \left\{ -\frac{\kappa^2}{2\sigma^2 t}(1+o(1)), \log \left(C \frac{t}{\kappa} \right) \right\} \sim \min \left\{ \frac{\kappa^2}{2\sigma^2 t}, \log \frac{\kappa}{t} \right\}.$$

It is easy to check that the asymptotic equality $\frac{\kappa^2}{2\sigma^2 t} \sim \log \frac{\kappa}{t}$ holds when $\kappa \sim \sigma \kappa_1(t)$. It follows that:

- in case $\kappa \leq \sigma \kappa_1(t)$ we have

$$-\log \frac{c(\kappa, t)}{\kappa} \sim \frac{\kappa^2}{2\sigma^2 t}; \quad (3.29)$$

- in case $\kappa \geq \sigma\kappa_1(t)$ we have

$$-\log \frac{c(\kappa, t)}{\kappa} \sim \log \frac{\kappa}{t}. \quad (3.30)$$

(Note that when $\kappa \sim \sigma\kappa_1(t)$ both relations (3.29) and (3.30) hold.)

We can deduce the asymptotic behavior of $\sigma_{\text{imp}}(\kappa, t)$ applying relation (2.38) (note that $c(\kappa, t)/\kappa \rightarrow 0$, since $\kappa \gg \sqrt{t}$) which, by Remark 2.10, reduces to the first line of (2.40), i.e.

$$\sigma_{\text{imp}}(\kappa, t) \sim \frac{\kappa}{\sqrt{2t(-\log(c(\kappa, t)/\kappa))}}. \quad (3.31)$$

Plugging (3.29)-(3.30) into this relation yields the first and second line of (3.27).

Next we assume that $t \rightarrow 0$ with $\eta \leq \kappa \ll \kappa_2(t)$, for some fixed $\eta > 0$. We claim that

$$-\log \frac{c(\kappa, t)}{\kappa} \sim \log \frac{1}{t}, \quad (3.32)$$

which plugged into (3.31) proves the second line of (3.27) (since $\log \frac{1}{t} \sim \log \frac{\kappa}{t}$ in this regime). When $\kappa > 0$ is fixed, (a stronger version of) relation (3.32) was proved by Figueroa-López and Forde in [FF12]. For the general case, we fix $a \in (0, \infty)$ and we apply relation (3.24), which yields the lower bound

$$\begin{aligned} c(\kappa, t) = \mathbb{E}[(e^{X_t} - e^\kappa)^+] &\geq (e^{a\kappa_2(t)} - e^\kappa) \mathbb{P}(X_t > a\kappa_2(t)) \sim e^{a\kappa_2(t) - f(a) \log \frac{1}{t}(1+o(1))} \\ &\sim e^{-f(a) \log \frac{1}{t}(1+o(1))}, \end{aligned} \quad (3.33)$$

because $\kappa_2(t) = \sqrt{\log \frac{1}{t}} \ll \log \frac{1}{t}$. For an upper bound, we recall that $\kappa \geq \eta$ and [FH09]

$$\log \mathbb{P}(X_t > \eta) \sim -\log \frac{1}{t} \quad \text{as } t \rightarrow 0. \quad (3.34)$$

Then, for every fixed $b \in (0, \infty)$, using (3.34), (3.24) and Cauchy-Schwarz we can write

$$\begin{aligned} c(\kappa, t) &\leq e^{b\kappa_2(t)} \mathbb{P}(\eta < X_t \leq b\kappa_2(t)) + \mathbb{E}[e^{X_t} \mathbf{1}_{\{X_t > b\kappa_2(t)\}}] \\ &\leq e^{b\kappa_2(t) - \log \frac{1}{t}(1+o(1))} + \mathbb{E}[e^{2X_t}]^{\frac{1}{2}} \mathbb{P}(X_t > b\kappa_2(t))^{\frac{1}{2}} \\ &\sim e^{-\log \frac{1}{t}(1+o(1))} + C_1 e^{-\frac{1}{2}f(b) \log \frac{1}{t}(1+o(1))}, \end{aligned}$$

where in the last step we used $\mathbb{E}[e^{2X_t}]^{\frac{1}{2}} \leq C_1$ for some constant C_1 , since $\mathbb{E}[e^{2X_t}] \rightarrow 1$ as $t \rightarrow 0$ by (3.20). By (3.23), we can fix b large enough so that $f(b) > 2$, hence $c(\kappa, t) \leq e^{-\log \frac{1}{t}(1+o(1))}$, and for every $\varepsilon > 0$ we can choose $a > 0$ small enough such that $f(a) < 1 + \varepsilon$, hence $c(\kappa, t) \geq e^{-(1+\varepsilon) \log \frac{1}{t}(1+o(1))}$ by (3.33). Altogether, relation (3.32) is proved.

Let us proceed with the regime $t \rightarrow 0$ and $\kappa \sim a\kappa_2(t)$ for some $a \in (0, \infty)$. Relation (3.24) shows that such a family of (κ, t) satisfies Hypothesis 2.2 with $I_+(\varrho) = f(\varrho a)/f(a)$ (we stress that $a \in (0, \infty)$ is fixed throughout this argument, hence I_+ can depend on a). Since the moment condition (2.9) is clearly satisfied by (3.20), we can apply Theorem 2.3: relation (2.18), coupled with (3.24), proves the third line of (3.27).

Finally, it remains to prove (3.28), hence we assume that either $t \rightarrow 0$ and $\kappa \gg \kappa_2(t)$, or $t \rightarrow \bar{t} \in (0, \infty)$ and $\kappa \rightarrow \infty$. Relation (3.25) shows that Hypothesis 2.2 holds with $I_+(\varrho) = \varrho$. By Theorem 2.3, relation (2.18) yields (3.28), completing the proof of Theorem 3.4. \square

3.3. The Heston Model. Given the parameters $\lambda, \vartheta, \eta, \sigma_0 \in (0, \infty)$ and $\varrho \in [-1, 1]$, the Heston model [H93] is a stochastic volatility model $(S_t)_{t \geq 0}$ defined by the following SDEs

$$\begin{cases} dS_t = S_t \sqrt{V_t} dW_t^1, \\ dV_t = -\lambda(V_t - \vartheta) dt + \eta \sqrt{V_t} dW_t^2, \\ X_0 = 0, \quad V_0 = \sigma_0, \end{cases}$$

where $(W_t^1)_{t \geq 0}$ and $(W_t^2)_{t \geq 0}$ are standard Brownian motions with $\langle dW_t^1, dW_t^2 \rangle = \varrho dt$.

Note that S_t displays explosion of moments, i.e. $E[S_T^p] = \infty$ for $p > 1$ large enough. In general, for any fixed $t \geq 0$ one can define the explosion moment $p^*(t)$ as

$$p^*(t) := \sup\{p > 0 : E[S_t^p] < \infty\},$$

so that $E[S_t^p] < \infty$ for $p < p^*(t)$ while $E[S_t^p] = \infty$ for $p > p^*(t)$ (in the case of Heston's model, one has $E[S_t^p] = \infty$ also for $p = p^*(t)$). The behavior of the explosion moment $p^*(t)$ is described in the following Lemma, proved below.

Lemma 3.5. *If $\varrho = -1$, then $p^*(t) = +\infty$ for every $t \geq 0$.*

If $\varrho > -1$, then $p^(t) \in (1, +\infty)$ for every $t > 0$. Moreover, as $t \downarrow 0$*

$$p^*(t) \sim \frac{C}{t},$$

where

$$C = C(\varrho, \eta) := \begin{cases} \frac{2}{\eta \sqrt{1 - \varrho^2}} \left(\arctan \frac{\sqrt{1 - \varrho^2}}{\varrho} + \pi \mathbf{1}_{\varrho < 0} \right) & \text{if } \varrho < 1 \\ \frac{2}{\eta} & \text{if } \varrho = 1 \end{cases}. \quad (3.35)$$

The asymptotic behavior of the implied volatility $\sigma_{\text{imp}}(\kappa, t)$ is known in the regimes of large strike (with fixed maturity) and small maturity (with fixed strike).

- In [BF08], Benaim and Friz show that for fixed $t > 0$, when $\kappa \rightarrow +\infty$

$$\sigma_{\text{imp}}(\kappa, t) \underset{\kappa \uparrow \infty}{\sim} \frac{\sqrt{2\kappa}}{\sqrt{t}} \left(\sqrt{p^*(t)} - \sqrt{p^*(t) - 1} \right), \quad (3.36)$$

based on the estimate (cf. also [AP07])

$$-\log P(X_t > \kappa) \underset{\kappa \uparrow \infty}{\sim} p^*(t) \kappa. \quad (3.37)$$

- In [FJ09], Forde and Jacquier have proved that for any fixed $\kappa > 0$, as $t \downarrow 0$

$$\sigma_{\text{imp}}(\kappa, t) \underset{t \downarrow 0}{\sim} \frac{\kappa}{\sqrt{2 \Lambda^*(\kappa)}}, \quad (3.38)$$

where $\Lambda^*(\cdot)$ is the Legendre transform of the function $\Lambda : \mathbb{R}^+ \rightarrow \mathbb{R}^+ \cup \{\infty\}$ given by

$$\Lambda(p) := \begin{cases} \frac{\sigma_0 p}{\eta \left(\sqrt{1 - \varrho^2} \cot \left(\frac{1}{2} \eta p \sqrt{1 - \varrho^2} \right) - \varrho \right)} & \text{if } p < C, \\ \infty & \text{if } p \geq C, \end{cases} \quad (3.39)$$

where C is the constant in (3.35). Their analysis is based on the estimate

$$-\log P(X_t \geq \kappa) \underset{t \downarrow 0}{\sim} \frac{1}{t} \Lambda^*(\kappa), \quad (3.40)$$

obtained by showing that the log-price $(X_t)_{t \geq 0}$ in the Heston model satisfies a large deviations principle as $t \downarrow 0$, with rate $1/t$ and good rate function $\Lambda^*(\kappa)$.

We first note that the asymptotics (3.36) and (3.38) follow easily from our Theorem 2.3, plugging the estimates (3.37) and (3.40) into relations (2.15) and (2.18), respectively.

We also observe that the estimates (3.36) and (3.38) match, in the following sense: if we take the limit $t \rightarrow 0$ of the right hand side of (3.36) (i.e. we first let $\kappa \uparrow +\infty$ and then $t \downarrow 0$ in $\sigma_{\text{imp}}(\kappa, t)$), we obtain

$$(3.36) \underset{t \downarrow 0}{\sim} \frac{\sqrt{2\kappa}}{\sqrt{t}} \frac{1}{2\sqrt{p^*(t)}} \sim \frac{\sqrt{2\kappa}}{\sqrt{t}} \frac{1}{2\sqrt{\frac{C}{t}}} = \frac{\sqrt{\kappa}}{\sqrt{2C}}. \quad (3.41)$$

If, on the other hand, we take the limit $\kappa \uparrow +\infty$ of the right hand side of (3.38) (i.e. we first let $t \downarrow 0$ and then $\kappa \uparrow +\infty$ in $\sigma_{\text{imp}}(\kappa, t)$), since $\Lambda^*(\kappa) \sim C\kappa$,[†] we obtain

$$(3.38) \underset{\kappa \uparrow +\infty}{\sim} \frac{\kappa}{\sqrt{2C\kappa}} = \frac{\sqrt{\kappa}}{\sqrt{2C}}, \quad (3.42)$$

which coincides with (3.41). Analogously, also the estimates (3.37) and (3.40) match.

It is then natural to conjecture that, for *any* family of values of (κ, t) such that $\kappa \uparrow +\infty$ and $t \downarrow 0$ jointly, one should have

$$\log P(X_t \geq \kappa) \sim -C \frac{\kappa}{t}, \quad (3.43)$$

where C is the constant in (3.35). If this holds, applying Theorem 2.3, relation (2.18) yields

$$\sigma_{\text{imp}}(\kappa, t) \sim \frac{\sqrt{\kappa}}{\sqrt{2C}}, \quad (3.44)$$

providing a smooth interpolation between (3.36) and (3.38).

Remark 3.6 (Surface asymptotics for the Heston model). If (3.44) holds for any family of values of (κ, t) with $\kappa \rightarrow \infty$ and $t \rightarrow 0$, it follows that for every $\varepsilon > 0$ there exists $M = M(\varepsilon) \in (0, \infty)$ such that the following inequalities hold:

$$(1 - \varepsilon) \frac{\sqrt{\kappa}}{\sqrt{2C}} \leq \sigma_{\text{imp}}(\kappa, t) \leq (1 + \varepsilon) \frac{\sqrt{\kappa}}{\sqrt{2C}},$$

for all (κ, t) in the region $\mathcal{A}_{T,M} := \{0 < t \leq \frac{1}{M}, \kappa > M\}$, as it follows easily by contradiction (cf. Remark 3.2 for a similar argument).

Proof of Lemma 3.5. Given any number $p > 1$ we define the explosion time $T^*(p)$ as

$$T^*(p) := \sup\{t > 0 : E[S_t^p] < \infty\}.$$

Note that if $T^*(p) = t \in (0, +\infty)$ then $p^*(t) = p$. By [AP07] (see also [FK09])

$$T^*(p) = \begin{cases} +\infty & \text{if } \Delta(p) \geq 0, \chi(p) < 0, \\ \frac{1}{\sqrt{\Delta(p)}} \log \left(\frac{\chi(p) + \sqrt{\Delta(p)}}{\chi(p) - \sqrt{\Delta(p)}} \right) & \text{if } \Delta(p) \geq 0, \chi(p) > 0, \\ \frac{2}{\sqrt{-\Delta(p)}} \left(\arctan \frac{\sqrt{-\Delta(p)}}{\chi(p)} + \pi 1_{\chi(p) < 0} \right) & \text{if } \Delta(p) < 0, \end{cases} \quad (3.45)$$

where

$$\chi(p) := \varrho \eta p - \lambda, \quad \Delta(p) := \chi^2(p) - \eta^2(p^2 - p),$$

[†]This is because $\Lambda(p) \uparrow +\infty$ as $p \uparrow C$, hence the slope of $\Lambda^*(\kappa)$ converges to C as $\kappa \rightarrow \infty$.

Observe that if $\varrho = -1$, then $\chi(p) = -\eta p - \lambda < 0$ and $\Delta(p) = \lambda^2 + p(2\eta\lambda + \eta^2) \geq 0$, which implies $T^*(p) = +\infty$ for every $p > 1$, or equivalently $p^*(t) = +\infty$ for every $t > 0$.

On the other hand, since

$$\Delta(p) = \varrho^2 \eta^2 p^2 + \lambda^2 - 2\eta\varrho\lambda p - \eta^2 p^2 + \eta^2 p = \eta^2 p^2 (\varrho^2 - 1) + p(\eta^2 - 2\eta\varrho\lambda) + \lambda^2,$$

we observe that if $\varrho \neq 1$, then $\Delta p < 0$ as $p \rightarrow +\infty$, which implies

$$\begin{aligned} T^*(p) &\underset{p \uparrow \infty}{\sim} \frac{2}{p(\eta\sqrt{1-\varrho^2})} \left(\arctan \frac{\eta p \sqrt{1-\varrho^2}}{\varrho\eta p} + \pi 1_{\varrho < 0} \right) \\ &= \frac{1}{p} \frac{2}{\eta\sqrt{1-\varrho^2}} \left(\arctan \frac{\sqrt{1-\varrho^2}}{\varrho} + \pi 1_{\varrho < 0} \right). \end{aligned} \quad (3.46)$$

In particular this leads to the conclusion that, if $|\varrho| \neq 1$, then

$$p^*(t) \underset{t \downarrow 0}{\sim} \frac{C}{t}$$

where C was defined in (3.35).

It remains to study the case $\varrho = 1$, in which $\chi(p) > 0$ for every p . We have two possibilities: if $\eta > 2\lambda$ then $\Delta(p) > 0$ when $p \rightarrow +\infty$, and so by (3.45)

$$T^*(p) \underset{p \uparrow \infty}{\sim} \frac{1}{\sqrt{p(\eta^2 + 2\eta\lambda)}} \log \left(1 + 2 \frac{\sqrt{p(\eta^2 + 2\eta\lambda)}}{\eta p - \sqrt{p(\eta^2 + 2\eta\lambda)}} \right) \sim \frac{2}{\eta} \frac{1}{p}.$$

On the other hand, if $\eta < 2\lambda$, then $\Delta(p) < 0$ when $p \rightarrow \infty$ and so

$$T^*(p) \underset{p \uparrow \infty}{\sim} \frac{2}{\sqrt{p(2\eta\lambda - \eta^2)}} \left(\arctan \frac{\sqrt{p(2\eta\lambda - \eta^2)}}{p\eta} \right) \sim \frac{2}{\eta} \frac{1}{p}.$$

Finally if $\eta = 2\lambda$, $\Delta(p) = \lambda^2$, and so

$$T^*(p) = \frac{1}{\lambda} \log \left(1 + \frac{2\lambda}{\eta p - 2\lambda} \right) \underset{p \uparrow \infty}{\sim} \frac{2}{\eta} \frac{1}{p}.$$

In all the cases we obtain $p^*(t) \underset{t \downarrow 0}{\sim} \frac{2}{\eta} \frac{1}{t}$, in agreement with (3.35). \square

4. FROM OPTION PRICE TO IMPLIED VOLATILITY

In this section we prove Theorem 2.9. We start with some background on Black&Scholes model and on related quantities. We let Z be a standard Gaussian random variable and denote by ϕ and Φ its density and distribution functions:

$$\phi(z) := \frac{P(Z \in dz)}{dz} = \frac{e^{-\frac{1}{2}z^2}}{\sqrt{2\pi}}, \quad \Phi(z) := P(Z \leq z) = \int_{-\infty}^z \phi(t) dt. \quad (4.1)$$

4.1. Mills ratio. The Mills ratio $U : \mathbb{R} \rightarrow (0, \infty)$ is defined by

$$U(z) := \frac{1 - \Phi(z)}{\phi(z)} = \frac{\Phi(-z)}{\phi(z)}, \quad \forall z \in \mathbb{R}. \quad (4.2)$$

The next lemma summarizes the main properties of U that will be used in the sequel.

Lemma 4.1. *The function U is smooth, strictly decreasing, strictly convex and satisfies*

$$U'(z) \sim -\frac{1}{z^2} \quad \text{as } z \uparrow \infty. \quad (4.3)$$

Proof. Since $\Phi'(z) = \phi(z)$ and ϕ is an analytic function, U is also analytic. Since $\phi'(z) = -z\phi(z)$, one obtains

$$U'(z) = zU(z) - 1, \quad U''(z) = U(z) + zU'(z) = (1 + z^2)U(z) - z. \quad (4.4)$$

Recalling that $U(z) > 0$, these relations already show that $U'(z) < 0$ and $U''(z) > 0$ for all $z \leq 0$. For $z > 0$, the following bounds hold [S54, eq. (19)], [P01, Th. 1.5]:

$$\frac{z}{z^2 + 1} = \frac{1}{z + \frac{1}{z}} < U(z) < \frac{1}{z + \frac{1}{z + \frac{2}{z}}} = \frac{z^2 + 2}{z^3 + 3z}, \quad \forall z > 0. \quad (4.5)$$

Applying (4.4) yields $U''(z) > 0$ and $-\frac{1}{1+z^2} < U'(z) < -\frac{1}{3+z^2}$ for all $z > 0$, hence (4.3). \square

We recall that the smooth function $D : (0, \infty) \rightarrow (0, \infty)$ was introduced in (2.32). Since

$$D'(z) = -\frac{1}{z^2}\phi(z) < 0, \quad (4.6)$$

$D(\cdot)$ is a strictly decreasing bijection (note that $\lim_{z \downarrow 0} D(z) = \infty$ and $\lim_{z \rightarrow \infty} D(z) = 0$). Its inverse $D^{-1} : (0, \infty) \rightarrow (0, \infty)$ is then smooth and strictly decreasing as well. Writing $D(z) = \phi(z)(\frac{1}{z} - U(z))$, it follows by (4.5) that $\frac{1}{z} - U(z) \sim \frac{1}{z^3}$ as $z \uparrow \infty$, hence

$$D(z) \sim \frac{1}{z^3}\phi(z) \sim \frac{e^{-\frac{1}{2}z^2}}{\sqrt{2\pi}z^3} \quad \text{as } z \uparrow \infty, \quad D(z) \sim \frac{1}{z}\phi(0) = \frac{1}{\sqrt{2\pi}z} \quad \text{as } z \downarrow 0.$$

It follows easily that $D^{-1}(\cdot)$ satisfies (2.33).

4.2. Black&Scholes. Let $(B_t)_{t \geq 0}$ be a standard Brownian motion. The Black&Scholes model is defined by a risk-neutral log-price $(X_t := \sigma B_t - \frac{1}{2}\sigma^2 t)_{t \geq 0}$, where the parameter $\sigma \in (0, \infty)$ represents the volatility. The Black&Scholes formula for the price of a normalized European call is $C_{BS}(\kappa, \sigma\sqrt{t})$, where κ is the log-strike, t is the maturity and we define

$$C_{BS}(\kappa, v) := E[(e^{vZ - \frac{1}{2}v^2} - e^\kappa)^+] = \begin{cases} (1 - e^\kappa)^+ & \text{if } v = 0, \\ \Phi(d_1) - e^\kappa\Phi(d_2) & \text{if } v > 0, \end{cases} \quad (4.7)$$

where Φ is defined in (4.1), and we set

$$\begin{cases} d_1 = d_1(\kappa, v) := -\frac{\kappa}{v} + \frac{v}{2}, \\ d_2 = d_2(\kappa, v) := -\frac{\kappa}{v} - \frac{v}{2}, \end{cases} \quad \text{so that} \quad \begin{cases} d_2 = d_1 - v, \\ d_2^2 = d_1^2 + 2\kappa. \end{cases} \quad (4.8)$$

Note that $C_{BS}(\kappa, v)$ is a continuous function of (κ, v) . Since $e^\kappa\phi(d_2) = \phi(d_1)$, for all $v > 0$ one easily computes

$$\frac{\partial C_{BS}(\kappa, v)}{\partial v} = \phi(d_1) > 0, \quad \frac{\partial C_{BS}(\kappa, v)}{\partial \kappa} = -e^\kappa\Phi(d_2) < 0,$$

hence $C_{BS}(\kappa, v)$ is strictly increasing in v and strictly decreasing in κ (see Figure 2). It is also directly checked that for all $\kappa \in \mathbb{R}$ and $v \geq 0$ one has

$$C_{BS}(\kappa, v) = 1 - e^\kappa + e^\kappa C_{BS}(-\kappa, v). \quad (4.9)$$

In the following key proposition, proved in Appendix A.4, we show that when $\kappa \geq 0$ the Black&Scholes call price $C_{BS}(\kappa, v)$ vanishes precisely when $v \rightarrow 0$ or $d_1 \rightarrow -\infty$ (or, more generally, in a combination of these two regimes, when $\min\{d_1, \log v\} \rightarrow -\infty$). We also provide sharp estimates on $C_{BS}(\kappa, v)$ for each regime (weaker estimates on $\log C_{BS}(\kappa, v)$ could be deduced from Theorems 2.3 and 2.4).

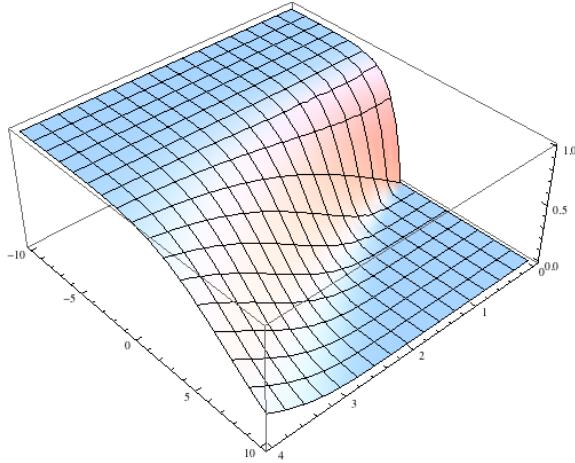


FIGURE 2. A plot of $(\kappa, v) \mapsto C_{BS}(\kappa, v)$, for $\kappa \in [-10, 10]$ and $v \in [0, 4]$.

Proposition 4.2. *For any family of values of (κ, v) with $\kappa \geq 0$, $v > 0$, one has*

$$C_{BS}(\kappa, v) \rightarrow 0 \quad \text{if and only if} \quad \min\{d_1, \log v\} \rightarrow -\infty, \quad (4.10)$$

that is, $C_{BS}(\kappa, v) \rightarrow 0$ if and only if from any subsequence of (κ, v) one can extract a sub-subsequence along which either $d_1 \rightarrow -\infty$ or $v \rightarrow 0$. Moreover:

- if $d_1 := -\frac{\kappa}{v} + \frac{v}{2} \rightarrow -\infty$, then

$$C_{BS}(\kappa, v) \sim \phi(d_1) \frac{v}{-d_1(-d_1 + v)}; \quad (4.11)$$

- if $v \rightarrow 0$, then

$$C_{BS}(\kappa, v) \sim -U'(-d_1) \phi(d_1) v; \quad (4.12)$$

where $\phi(\cdot)$ and $U(\cdot)$ are defined in (4.1) and (4.2).

4.3. Proof of Theorem 2.9. Since the function $v \mapsto C_{BS}(\kappa, v)$ is a bijection from $[0, \infty)$ to $[(1 - e^\kappa)^+, 1]$, it admits an inverse function $c \mapsto V_{BS}(\kappa, c)$, defined by

$$C_{BS}(\kappa, V_{BS}(\kappa, c)) = c. \quad (4.13)$$

By construction, $V_{BS}(\kappa, \cdot)$ is a strictly increasing bijection from $[(1 - e^\kappa)^+, 1]$ to $[0, \infty)$. We will mainly focus on the case $\kappa \geq 0$, for which $V_{BS}(\kappa, \cdot) : [0, 1] \rightarrow [0, \infty)$.

Consider an arbitrary model, with a risk-neutral log-price $(X_t)_{t \geq 0}$, and let $c(\kappa, t)$ be the corresponding price of a normalized European call option, cf. (2.1). Since $z \mapsto (z - e^\kappa)^+$ is a convex function, one has $c(\kappa, t) \geq (E[e^{X_t}] - e^\kappa)^+ = (1 - e^\kappa)^+$ by Jensen's inequality; since $(z - e^\kappa)^+ < z^+$, one has $c(\kappa, t) < E[e^{X_t}] = 1$. Consequently, by (4.13), we have the following relation between the *implied volatility* $\sigma_{\text{imp}}(\kappa, t)$ (defined in §2.1) and $V_{BS}(\kappa, c(\kappa, t))$:

$$\sigma_{\text{imp}}(\kappa, t) := \frac{V_{BS}(\kappa, c(\kappa, t))}{\sqrt{t}}. \quad (4.14)$$

Relation (4.14) allows to reformulate Theorem 2.9 more transparently in terms of the function V_{BS} . Inspired by (2.2), we define $p = p(\kappa, c)$ by

$$p := c - (1 - e^\kappa). \quad (4.15)$$

Consider an arbitrary family of values of (κ, c) , such that either $\kappa \geq 0$, $c \in (0, 1)$ and $c \rightarrow 0$, or alternatively $\kappa \leq 0$, $p \in (0, 1)$ and $p \rightarrow 0$ (with p as in (4.15)). Then, in light of (4.14), we can write the following:

- If κ bounded away from zero ($\liminf |\kappa| > 0$), relation (2.37) is equivalent to

$$V_{BS}(\kappa, c) \sim \begin{cases} \sqrt{2(-\log c + \kappa)} - \sqrt{2(-\log c)} & \text{if } \kappa > 0, \\ \sqrt{2(-\log p)} - \sqrt{2(-\log p + \kappa)} & \text{if } \kappa < 0. \end{cases} \quad (4.16)$$

- If κ is bounded away from infinity ($\limsup |\kappa| < \infty$), relations (2.38) and (2.39) are equivalent to

$$V_{BS}(\kappa, c) \sim \begin{cases} \frac{\kappa}{D^{-1}(\frac{c}{\kappa})} & \text{if } \kappa > 0, \\ \sqrt{2\pi} c = \sqrt{2\pi} p & \text{if } \kappa = 0, \\ \frac{-\kappa}{D^{-1}(\frac{p}{-\kappa})} & \text{if } \kappa < 0, \end{cases} \quad (4.17)$$

where $D^{-1}(\cdot)$ is the inverse of the function $D(\cdot)$ defined in (2.32), and satisfies (2.33).

The proof of Theorem 2.9 is now reduced to proving relations (4.16) and (4.17). We first show that we can assume $\kappa \geq 0$, by a symmetry argument.

Deducing the case $\kappa \leq 0$ from the case $\kappa \geq 0$. Recalling (4.9) and (4.13), for all $\kappa \in \mathbb{R}$ and $c \in [(1 - e^\kappa)^+, 1)$ we have

$$V_{BS}(\kappa, c) = V_{BS}(-\kappa, 1 - e^{-\kappa} + e^{-\kappa}c) = V_{BS}(-\kappa, e^{-\kappa}p),$$

where p is defined in (4.15). As a consequence, in the case $\kappa \leq 0$, replacing κ by $-\kappa$ and c by $e^{-\kappa}p$ in the first line of (4.16), one obtains the second line of (4.16).

Performing the same replacements in the first line of (4.17) yields

$$V_{BS}(\kappa, c) \sim \frac{-\kappa}{D^{-1}(e^{-\kappa} \frac{p}{-\kappa})},$$

which is slightly different with respect to the third line of (4.17). However, the discrepancy is only apparent, because we claim that $D^{-1}(e^{-\kappa} \frac{p}{-\kappa}) \sim D^{-1}(\frac{p}{-\kappa})$. This is checked as follows: if $\kappa \rightarrow 0$, then $e^{-\kappa} \frac{p}{-\kappa} \sim \frac{p}{-\kappa}$; if $\kappa \rightarrow \bar{\kappa} \in (-\infty, 0)$, since $p \rightarrow 0$ by assumption, the first relation in (2.33) yields $D^{-1}(e^{-\kappa} \frac{p}{-\kappa}) \sim \sqrt{2(-\log(\frac{p}{-\kappa}) + \bar{\kappa})} \sim \sqrt{2(-\log(\frac{p}{-\bar{\kappa}}))} \sim D^{-1}(\frac{p}{-\bar{\kappa}})$, as required. (See the lines following (4.26) below for more details.) \square

Proof of (4.16) for $\kappa \geq 0$. We fix a family of values of (κ, c) with $c \rightarrow 0$ and κ bounded away from zero, say $\kappa \geq \delta$ for some fixed $\delta > 0$. Our goal is to prove that relation (4.16) holds. If we set $v := V_{BS}(\kappa, c)$, by definition (4.13) we have $C_{BS}(\kappa, v) = c \rightarrow 0$.

Let us first show that $d_1 := -\frac{\kappa}{v} + \frac{v}{2} \rightarrow -\infty$. By Proposition 4.2, $C_{BS}(\kappa, v) \rightarrow 0$ implies $\min\{d_1, \log v\} \rightarrow -\infty$, which means that every subsequence of values of (κ, c) admits a further sub-subsequence along which either $d_1 \rightarrow \infty$ or $v \rightarrow 0$. The key point is that $v \rightarrow 0$ implies $d_1 \rightarrow -\infty$, because $d_1 \leq -\frac{\delta}{v} + \frac{v}{2}$ (recall that $\kappa \geq \delta$). Thus $d_1 \rightarrow -\infty$ along every sub-subsequence, which means that $d_1 \rightarrow -\infty$ along the whole family of values of (κ, c) .

Since $d_1 \rightarrow -\infty$, we can apply relation (4.11). Taking log of both sides of that relation, recalling the definition (4.1) of ϕ and the fact that $C_{BS}(\kappa, v) = c$, we can write

$$\log c \sim -\frac{1}{2}d_1^2 - \log \sqrt{2\pi} + \log \frac{v}{-d_1(-d_1 + v)}. \quad (4.18)$$

We now show that the last term in the right hand side is $o(d_1^2)$ and can therefore be neglected. Note that $-d_1 \geq 1$ eventually, because $d_1 \rightarrow -\infty$, hence

$$\log \frac{v}{-d_1(-d_1+v)} \leq \log \frac{v}{1+v} \leq 0.$$

Since $v \mapsto \frac{-d_1+v}{v}$ is decreasing for $-d_1 > 0$, in case $v \geq -d_1$ one has

$$\left| \log \frac{v}{-d_1(-d_1+v)} \right| = \log \frac{-d_1(-d_1+v)}{v} \leq \log(-2d_1) = o(d_1^2).$$

On the other hand, recalling that $d_1 \leq -\frac{\delta}{v} + \frac{v}{2}$, in case $v < -d_1$ one has $d_1 \leq -\frac{\delta}{v} - \frac{d_1}{2}$, which can be rewritten as $v \geq \frac{2\delta}{-3d_1}$ and together with $v < -d_1$ yields

$$\left| \log \frac{v}{-d_1(-d_1+v)} \right| = \log \frac{-d_1(-d_1+v)}{v} \leq \log \frac{-d_1(-d_1-d_1)}{\frac{2\delta}{-3d_1}} = \log \left(\frac{3(-d_1)^3}{2\delta} \right) = o(d_1^2).$$

In conclusion, (4.18) yields $\log c \sim -\frac{1}{2}d_1^2$, that is there exists $\gamma = \gamma(\kappa, c) \rightarrow 0$ such that $(1+\gamma)\log c = -\frac{1}{2}d_1^2$, and since $\log c \leq 0$ we can write

$$(1+\gamma)|\log c| = \frac{1}{2}d_1^2 = \frac{1}{2}\left(\frac{\kappa^2}{v^2} + \frac{v^2}{4} - \kappa\right).$$

This is a second degree equation in v^2 , whose solutions (both positive) are

$$v^2 = 2\kappa \left[1 + 2\frac{(1+\gamma)|\log c|}{\kappa} \pm 2\sqrt{\left(\frac{(1+\gamma)|\log c|}{\kappa}\right)^2 + \frac{(1+\gamma)|\log c|}{\kappa}} \right]. \quad (4.19)$$

Since $d_1 \rightarrow -\infty$, eventually one has $d_1 < 0$: since $d_1 = -\frac{\kappa}{v} + \frac{v}{2} = -\frac{1}{2v}(\sqrt{2\kappa} - v)(\sqrt{2\kappa} + v)$, it follows that $v^2 < 2\kappa$, which selects the “-” solution in (4.19). Taking square roots of both sides of (4.19) and recalling that $v = V_{\text{BS}}(\kappa, c)$ yields the equality

$$V_{\text{BS}}(\kappa, c) = \sqrt{2(1+\gamma)|\log c| + 2\kappa} - \sqrt{2(1+\gamma)|\log c|}, \quad (4.20)$$

as one checks squaring both sides of (4.20).

Finally, since $\gamma \rightarrow 0$, it is quite intuitive that relation (4.20) yields (4.16). To prove this fact, we observe that by (4.20) we can write

$$\frac{V_{\text{BS}}(\kappa, c)}{\sqrt{2|\log c| + 2\kappa} - \sqrt{2|\log c|}} = f_\gamma\left(\frac{\kappa}{|\log c|}\right), \quad (4.21)$$

where for fixed $\gamma > -1$ we define the function $f_\gamma : [0, \infty) \rightarrow (0, \infty)$ by

$$f_\gamma(x) := \frac{\sqrt{1+\gamma+x} - \sqrt{1+\gamma}}{\sqrt{1+x} - 1} \quad \text{for } x > 0, \quad f_\gamma(0) := \lim_{x \downarrow 0} f_\gamma(x) = \frac{1}{\sqrt{1+\gamma}}.$$

By direct computation, when $\gamma > 0$ (resp. $\gamma < 0$) one has $\frac{d}{dx}f_\gamma(x) > 0$ (resp. < 0) for all $x > 0$. Since $\lim_{x \rightarrow +\infty} f_\gamma(x) = 1$, it follows that for every $x \geq 0$ one has $f_\gamma(0) \leq f_\gamma(x) \leq 1$ if $\gamma > 0$, while $1 \leq f_\gamma(x) \leq f_\gamma(0)$ if $\gamma < 0$; consequently, for any γ ,

$$\frac{1}{\sqrt{1+|\gamma|}} \leq f_\gamma(x) \leq \frac{1}{\sqrt{1-|\gamma|}}, \quad \forall x \geq 0,$$

which yields $\lim_{\gamma \rightarrow 0} f_\gamma(x) = 1$ uniformly over $x \geq 0$. By (4.21), relation (4.16) is proved. \square

Proof of (4.17) for $\kappa \geq 0$. We now fix a family of values of (κ, c) with $c \rightarrow 0$ and κ bounded away from infinity, say $0 \leq \kappa \leq M$ for some fixed $M \in (0, \infty)$, and we prove relation (4.17).

We set $v := V_{\text{BS}}(\kappa, c)$ so that $C_{\text{BS}}(\kappa, v) = c \rightarrow 0$, cf. (4.13). (Note that $v > 0$, because $c > 0$ by assumption.) Applying Proposition 4.2 we have $\min\{d_1, \log v\} \rightarrow -\infty$, i.e. either $d_1 \rightarrow -\infty$ or $v \rightarrow 0$ along sub-subsequences. However, this time $d_1 \rightarrow -\infty$ implies $v \rightarrow 0$, because $d_1 \geq -\frac{M}{v} + \frac{v}{2}$ (recall that $\kappa \leq M$), which means that $v \rightarrow 0$ along the whole given family of values of (κ, c) .

Since $v \rightarrow 0$, relation (4.12) yields

$$c \sim -U'(-d_1) \phi(d_1) v. \quad (4.22)$$

Let us focus on $U'(-d_1)$: recalling that $d_1 = -\frac{\kappa}{v} + \frac{v}{2}$ and $v \rightarrow 0$, we first show that

$$U'(-d_1) \sim U'\left(\frac{\kappa}{v}\right). \quad (4.23)$$

By a subsequence argument, we may assume that $\frac{\kappa}{v} \rightarrow \varrho \in [0, \infty]$, and we recall that $v \rightarrow 0$:

- if $\varrho < \infty$, $U'(-d_1)$ and $U'\left(\frac{\kappa}{v}\right)$ converge to $U'(\varrho) \neq 0$, hence $U'(-d_1)/U'\left(\frac{\kappa}{v}\right) \rightarrow 1$;
- if $\varrho = \infty$, $-d_1$ and $\frac{\kappa}{v}$ diverge to ∞ and (4.3) yields $U'(-d_1)/U'\left(\frac{\kappa}{v}\right) \sim (\frac{\kappa}{v})/(-d_1) \rightarrow 1$.

The proof of (4.23) is completed. Next we observe that, again by $v \rightarrow 0$,

$$\phi(-d_1) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}d_1^2} = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{\kappa^2}{v^2} + \frac{v^2}{2} - \kappa)} \sim e^{\frac{1}{2}\kappa} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\frac{\kappa^2}{v^2}} = e^{\frac{1}{2}\kappa} \phi\left(\frac{\kappa}{v}\right).$$

We can thus rewrite (4.22) as

$$c \sim -U'\left(\frac{\kappa}{v}\right) \phi\left(\frac{\kappa}{v}\right) e^{\frac{1}{2}\kappa} v. \quad (4.24)$$

If $\kappa = 0$, recalling (4.4) we obtain $c \sim \phi(0)v = \frac{1}{\sqrt{2\pi}}v$, which is the second line of (4.17).

Next we assume $\kappa > 0$. By (4.4), (4.2) and (2.32), for all $z > 0$ we can write

$$-U'(z) \phi(z) = -\phi(z)(zU(z) - 1) = \phi(z) - z\Phi(-z) = zD(z),$$

hence (4.24) can be rewritten as

$$c \sim \kappa e^{\frac{1}{2}\kappa} D\left(\frac{\kappa}{v}\right), \quad \text{i.e.} \quad (1 + \gamma)c = \kappa e^{\frac{1}{2}\kappa} D\left(\frac{\kappa}{v}\right),$$

for some $\gamma = \gamma(\kappa, c) \rightarrow 0$. Recalling that $v = V_{\text{BS}}(\kappa, c)$, we have shown that

$$V_{\text{BS}}(\kappa, c) = \frac{\kappa}{D^{-1}\left(\frac{(1+\gamma)c}{\kappa e^{\frac{1}{2}\kappa}}\right)}. \quad (4.25)$$

We now claim that

$$D^{-1}\left(\frac{(1+\gamma)c}{\kappa e^{\frac{1}{2}\kappa}}\right) \sim D^{-1}\left(\frac{c}{\kappa}\right). \quad (4.26)$$

By a subsequence argument, we may assume that $\frac{c}{\kappa} \rightarrow \eta \in [0, \infty]$ and $\kappa \rightarrow \bar{\kappa} \in [0, M]$.

- If $\eta \in (0, \infty)$, then $\bar{\kappa} = 0$ (recall that $c \rightarrow 0$) hence $(1 + \gamma)c/(\kappa e^{\frac{1}{2}\kappa}) \rightarrow \eta$; then both sides of (4.26) converge to $D^{-1}(\eta) \in (0, \infty)$, hence their ratio converges to 1.
- If $\eta = \infty$, then again $\bar{\kappa} = 0$, hence $(1 + \gamma)c/(\kappa e^{\frac{1}{2}\kappa}) \rightarrow \infty$: since $D^{-1}(y) \sim \frac{1}{\sqrt{2\pi}}y^{-1}$ as $y \rightarrow \infty$, cf. (2.33), it follows immediately that (4.26) holds.

- If $\eta = 0$, then $(1 + \gamma)c/(\kappa e^{\frac{1}{2}\kappa}) \rightarrow 0$: since $D^{-1}(y) \sim \sqrt{2|\log y|}$ as $y \rightarrow 0$, cf. (2.33),

$$D^{-1}\left(\frac{(1 + \gamma)c}{\kappa e^{\frac{1}{2}\kappa}}\right) \sim \sqrt{2\left|\left(\log \frac{c}{\kappa}\right) + \left(\log \frac{1 + \gamma}{e^{\frac{1}{2}\kappa}}\right)\right|} \sim \sqrt{2\left|\log \frac{c}{\kappa}\right|},$$

because $|\log \frac{c}{\kappa}| \rightarrow \infty$ while $|\log[(1 + \gamma)/e^{\frac{1}{2}\kappa}]| \rightarrow \frac{1}{2}\bar{\kappa} \in [0, \frac{M}{2}]$, hence (4.26) holds.

Having proved (4.26), we can plug it into (4.25), obtaining precisely the first line of (4.17). This completes the proof of Theorem 2.9. \square

5. FROM TAIL PROBABILITY TO OPTION PRICE

In this section we prove Theorems 2.3, 2.4 and 2.7. We stress that it is enough to prove the asymptotic relations for the option prices $c(\kappa, t)$ and $p(-\kappa, t)$, because the corresponding relations for the implied volatility $\sigma_{\text{imp}}(\pm\kappa, t)$ follow immediately applying Theorem 2.9.

5.1. Proof of Theorem 2.3 and 2.4. We prove Theorem 2.3 and 2.4 at the same time. We recall that the tail probabilities $\bar{F}_t(\kappa)$, $F_t(-\kappa)$ are defined in (1.1). Throughout the proof, we fix a family of values of (κ, t) with $\kappa > 0$ and $0 < t < T$, for some fixed $T \in (0, \infty)$, such that Hypothesis 2.2 is satisfied.

Extracting subsequences, we may distinguish three regimes for κ :

- if $\kappa \rightarrow \infty$ our goal is to prove (2.14), resp. (2.21);
- if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$ our goal is to prove (2.17), resp. (2.23), because in this case, plainly, one has $-\log \bar{F}_t(\kappa)/\kappa \rightarrow \infty$, resp. $-\log F_t(-\kappa)/\kappa \rightarrow \infty$, by (2.5);
- if $\kappa \rightarrow 0$, our goal is to prove (2.19), resp. (2.26).

Of course, each regime has different assumptions, as in Theorem 2.3 and 2.4.

Step 0. Preparation. It follows by conditions (2.7) and (2.8) that

$$\forall \varepsilon > 0 \quad \exists \varrho_\varepsilon \in (1, \infty) : \quad I_\pm(\varrho_\varepsilon) < 1 + \varepsilon, \quad (5.1)$$

therefore for every $\varepsilon > 0$ one has eventually

$$\begin{aligned} \log \bar{F}_t(\varrho_\varepsilon \kappa) &\geq (1 + \varepsilon) \log \bar{F}_t(\kappa), \quad \text{resp.} \\ \log F_t(-\varrho_\varepsilon \kappa) &\geq (1 + \varepsilon) \log F_t(-\kappa), \end{aligned} \quad (5.2)$$

where the inequality is “ \geq ” instead of “ \leq ”, because both sides are negative quantities.

We stress that $\bar{F}_t(\kappa) \rightarrow 0$, resp. $F_t(-\kappa) \rightarrow 0$, by (2.5), hence

$$\log \bar{F}_t(\kappa) \rightarrow -\infty, \quad \text{resp.} \quad \log F_t(-\kappa) \rightarrow -\infty. \quad (5.3)$$

Moreover, we claim that in any of the regimes $\kappa \rightarrow \infty$, $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$ and $\kappa \rightarrow 0$ one has

$$\log \bar{F}_t(\kappa) + \kappa \rightarrow -\infty. \quad (5.4)$$

This follows readily by (5.3) if $\kappa \rightarrow 0$ or $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$. If $\kappa \rightarrow \infty$ we argue as follows: by Markov's inequality, for $\eta > 0$

$$\bar{F}_t(\kappa) \leq \mathbb{E}[e^{(1+\eta)X_t}]e^{-(1+\eta)\kappa}, \quad (5.5)$$

hence

$$\log \bar{F}_t(\kappa) + \kappa \leq -\eta\kappa + \log \mathbb{E}[e^{(1+\eta)X_t}].$$

Since in the regime $\kappa \rightarrow \infty$ we assume that the moment condition (2.9) holds for some or every $\eta > 0$, the term $\log E[e^{(1+\eta)X_t}]$ is bounded from above, hence eventually

$$\log \bar{F}_t(\kappa) + \kappa \leq -\frac{\eta}{2} \kappa, \quad (5.6)$$

which proves relation (5.4).

The rest of the proof is divided in four steps, in each of which we prove lower and upper bounds on $c(\kappa, t)$ and $p(-\kappa, t)$, respectively.

Step 1. Lower bounds on $c(\kappa, t)$. We are going to prove sharp lower bounds on $c(\kappa, t)$, that will lead to relations (2.14), (2.17) and (2.19).

By (2.1) and (5.1), for every $\varepsilon > 0$ we can write

$$c(\kappa, t) \geq E[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \varrho_\varepsilon \kappa\}}] \geq (e^{\varrho_\varepsilon \kappa} - e^\kappa) \bar{F}_t(\varrho_\varepsilon \kappa), \quad (5.7)$$

and applying (5.2) we get

$$\log c(\kappa, t) \geq \log(e^{\varrho_\varepsilon \kappa} - e^\kappa) + (1 + \varepsilon) \log \bar{F}_t(\kappa). \quad (5.8)$$

If $\kappa \rightarrow \infty$, since $\log(e^{\varrho_\varepsilon \kappa} - e^\kappa) = \kappa + \log(e^{(\varrho_\varepsilon - 1)\kappa} - 1) \geq \kappa$ eventually, we obtain

$$\begin{aligned} \log c(\kappa, t) &\geq \kappa + (1 + \varepsilon) \log \bar{F}_t(\kappa) = (1 + \varepsilon)(\log \bar{F}_t(\kappa) + \kappa) - \varepsilon \kappa \\ &\geq (1 + \varepsilon + \frac{2}{\eta} \varepsilon)(\log \bar{F}_t(\kappa) + \kappa), \end{aligned} \quad (5.9)$$

where in the last inequality we have applied (5.6). It follows that

$$\limsup \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa) + \kappa} \leq 1 + \varepsilon + \frac{2}{\eta} \varepsilon, \quad (5.10)$$

where the lim sup is taken along the given family of values of (κ, t) (note that $\log c(\kappa, t)$ and $\log \bar{F}_t(\kappa) + \kappa$ are negative quantities, cf. (5.4), hence the reverse inequality with respect to (5.9)). Since $\varepsilon > 0$ is arbitrary and $\eta > 0$ is fixed, we have shown that

$$\limsup \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa) + \kappa} \leq 1, \quad (5.11)$$

that is we have obtained a sharp bound for (2.14).

If $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$, since $\log(e^{\varrho_\varepsilon \kappa} - e^\kappa) \rightarrow \log(e^{\varrho_\varepsilon \bar{\kappa}} - e^{\bar{\kappa}})$ is bounded while $\log \bar{F}_t(\kappa) \rightarrow -\infty$, relation (5.8) gives

$$\limsup \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa)} \leq 1 + \varepsilon.$$

Since $\varepsilon > 0$ is arbitrary, we have shown that when $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$

$$\limsup \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa)} \leq 1, \quad (5.12)$$

obtaining a sharp bound for (2.17).

Finally, if $\kappa \rightarrow 0$, since for $\kappa \geq 0$ by convexity $\log(e^{\varrho_\varepsilon \kappa} - e^\kappa) = \kappa + \log(e^{(\varrho_\varepsilon - 1)\kappa} - 1) \geq \kappa + \log((\varrho_\varepsilon - 1)\kappa) = \kappa + \log(\varrho_\varepsilon - 1) + \log \kappa$, relation (5.8) yields

$$\log \frac{c(\kappa, t)}{\kappa} = \log c(\kappa, t) - \log \kappa \geq \log(\varrho_\varepsilon - 1) + (1 + \varepsilon) \log \bar{F}_t(\kappa).$$

Again, since $\log(\varrho_\varepsilon - 1)$ is constant and $\log \bar{F}_t(\kappa) \rightarrow -\infty$, and $\varepsilon > 0$ is arbitrary, we get

$$\limsup \frac{\log(c(\kappa, t)/\kappa)}{\log \bar{F}_t(\kappa)} \leq 1, \quad (5.13)$$

proving a sharp bound for (2.19).

Step 2. Lower bounds on $p(-\kappa, t)$. We are going to prove sharp lower bounds on $p(-\kappa, t)$, that will lead to relations (2.21), (2.23) and (2.26).

Recalling (2.1) and (5.1), for every $\varepsilon > 0$ we can write

$$p(-\kappa, t) \geq \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t \leq -\varrho_\varepsilon \kappa\}}] \geq (e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) F_t(-\varrho_\varepsilon \kappa), \quad (5.14)$$

and applying (5.2) we obtain

$$\log p(-\kappa, t) \geq \log(e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) + (1 + \varepsilon) \log F_t(-\kappa). \quad (5.15)$$

If $\kappa \rightarrow \infty$, since $\log(e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) = -\kappa + \log(1 - e^{-(\varrho_\varepsilon - 1)\kappa}) \sim -\kappa$, eventually one has $\log(e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) \geq -(1 + \varepsilon)\kappa$ and we obtain

$$\log p(-\kappa, t) \geq (1 + \varepsilon)(\log F_t(-\kappa) - \kappa).$$

Since $\varepsilon > 0$ is arbitrary, it follows that

$$\limsup \frac{\log p(-\kappa, t)}{\log F_t(-\kappa) - \kappa} \leq 1, \quad (5.16)$$

which is a sharp bound for (2.21).

If $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$, since $\log(e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) \rightarrow \log(e^{-\bar{\kappa}} - e^{-\varrho_\varepsilon \bar{\kappa}})$ is bounded while $\log F_t(-\kappa) \rightarrow -\infty$, and $\varepsilon > 0$ is arbitrary, relation (5.15) gives

$$\limsup \frac{\log p(-\kappa, t)}{\log F_t(-\kappa)} \leq 1, \quad (5.17)$$

which is a sharp bound for (2.23).

Finally, if $\kappa \rightarrow 0$, since $e^{-\kappa} - e^{-\varrho_\varepsilon \kappa} = e^{-\varrho_\varepsilon \kappa}(e^{(\varrho_\varepsilon - 1)\kappa} - 1) \geq e^{-\varrho_\varepsilon \kappa}(\varrho_\varepsilon - 1)\kappa$ by convexity, since $\kappa \geq 0$, one has eventually

$$\log(e^{-\kappa} - e^{-\varrho_\varepsilon \kappa}) \geq \log \kappa + \log(e^{-\varrho_\varepsilon \kappa}(\varrho_\varepsilon - 1)) \geq \log \kappa + \varepsilon \log F_t(-\kappa),$$

because $\log(e^{-\varrho_\varepsilon \kappa}(\varrho_\varepsilon - 1)) \rightarrow \log(\varrho_\varepsilon - 1) > -\infty$ while $\log F_t(-\kappa) \rightarrow -\infty$. Relation (5.15) then yields, eventually,

$$\log \frac{p(-\kappa, t)}{\kappa} = \log p(-\kappa, t) - \log \kappa \geq (1 + 2\varepsilon) \log F_t(-\kappa).$$

Since $\varepsilon > 0$ is arbitrary, we have shown that

$$\limsup \frac{\log(p(-\kappa, t)/\kappa)}{\log F_t(-\kappa)} \leq 1, \quad (5.18)$$

obtaining a sharp bound for (2.26).

Step 3. Upper bounds on $c(\kappa, t)$. We are going to prove sharp upper bounds on $c(\kappa, t)$, that will complete the proof of relations (2.14), (2.17) and (2.19). We first consider the case when *the moment assumptions* (2.9) and (2.11) hold for every $\eta > 0$.

Let us look at the regimes $\kappa \rightarrow \infty$ and $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$ (i.e. κ is bounded away from zero), assuming that condition (2.9) holds for every $\eta > 0$. By Hölder's inequality,

$$c(\kappa, t) = \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa\}}] \leq \mathbb{E}[e^{X_t} \mathbf{1}_{\{X_t > \kappa\}}] \leq \mathbb{E}[e^{(1+\eta)X_t}]^{\frac{1}{1+\eta}} \bar{F}_t(\kappa)^{\frac{\eta}{1+\eta}}. \quad (5.19)$$

Let us fix $\varepsilon > 0$ and choose $\eta = \eta_\varepsilon$ large enough, so that $\frac{\eta}{1+\eta} > 1 - \varepsilon$. By assumption (2.9), for some $C \in (0, \infty)$ one has

$$\mathbb{E}[e^{(1+\eta)X_t}]^{\frac{1}{1+\eta}} \leq C,$$

hence eventually, recalling that $\log \bar{F}_t(\kappa) \rightarrow -\infty$, by (5.3),

$$\log c(\kappa, t) \leq \log C + (1 - \varepsilon) \log \bar{F}_t(\kappa) \leq (1 - 2\varepsilon) \log \bar{F}_t(\kappa). \quad (5.20)$$

Since $\varepsilon > 0$ is arbitrary, this shows that

$$\liminf \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa)} \geq 1. \quad (5.21)$$

which together with (5.12) completes the proof of (2.17), if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$. If $\kappa \rightarrow \infty$ and condition (2.9) holds for every $\eta > 0$, then $\log \bar{F}_t(\kappa)/\kappa \rightarrow -\infty$ by (5.5), which yields $\log \bar{F}_t(\kappa) \sim \log \bar{F}_t(\kappa) + \kappa$, hence (5.21) together with (5.11) completes the proof of (2.14).

We then consider the regime $\kappa \rightarrow 0$, assuming that condition (2.11) holds for every $\eta > 0$. We modify (5.19) as follows: since $(e^{X_t} - e^\kappa) \leq (e^{X_t} - 1) \leq |e^{X_t} - 1|$,

$$c(\kappa, t) \leq \mathbb{E}[|e^{X_t} - 1| \mathbf{1}_{\{X_t > \kappa\}}] \leq \kappa \mathbb{E}\left[\left|\frac{e^{X_t} - 1}{\kappa}\right|^{1+\eta}\right]^{\frac{1}{1+\eta}} \bar{F}_t(\kappa)^{\frac{\eta}{1+\eta}}. \quad (5.22)$$

Let us fix $\varepsilon > 0$ and choose $\eta = \eta_\varepsilon$ large enough, so that $\frac{\eta}{1+\eta} > 1 - \varepsilon$. By assumption (2.11), for some $C \in (0, \infty)$ one has

$$\mathbb{E}\left[\left|\frac{e^{X_t} - 1}{\kappa}\right|^{1+\eta}\right]^{\frac{1}{1+\eta}} \leq C, \quad (5.23)$$

hence relation (5.22) yields eventually

$$\log \frac{c(\kappa, t)}{\kappa} \leq \log C + (1 - \varepsilon) \log \bar{F}_t(\kappa) \leq (1 - 2\varepsilon) \log \bar{F}_t(\kappa). \quad (5.24)$$

Since $\varepsilon > 0$ is arbitrary, we have proved that

$$\liminf \frac{\log (c(\kappa, t)/\kappa)}{\log \bar{F}_t(\kappa)} \geq 1, \quad (5.25)$$

which together with (5.13) completes the proof of (2.19).

It remains to consider the case when the moment assumptions (2.9) and (2.11) holds for some $\eta > 0$, but in addition conditions (2.13) (if $\kappa \rightarrow \infty$ or $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$) or (2.16) (if $\kappa \rightarrow 0$) holds. We start with considerations that are valid in any regime of κ .

Defining the constant

$$A := \limsup \left\{ \frac{-\kappa}{\log \bar{F}_t(\kappa) + \kappa} \right\} + 1, \quad (5.26)$$

where the \limsup is taken along the given family of values of (κ, t) , we claim that $A < \infty$. This follows by (5.4) if $\kappa \rightarrow 0$ or if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$ (in which case, plainly, $A = 1$), while if $\kappa \rightarrow +\infty$ it suffices to apply (5.6) to get $A \leq 2/\eta + 1$. It follows by (5.26) that eventually

$$\kappa \leq -A(\log \bar{F}_t(\kappa) + \kappa). \quad (5.27)$$

Next we show that, for all fixed $\varepsilon > 0$ and $1 < M < \infty$, eventually one has

$$\log \left(\sup_{y \in [1, M]} e^{\kappa y} \bar{F}_t(\kappa y) \right) \leq (1 - \varepsilon) (\log \bar{F}_t(\kappa) + \kappa), \quad (5.28)$$

which means that the sup is approximately attained for $y = 1$. This is easy if $\kappa \rightarrow 0$ or if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$: in fact, since $\kappa \rightarrow \bar{F}_t(\kappa)$ is non-increasing, we can write

$$\begin{aligned} \log \left(\sup_{y \in [1, M]} e^{\kappa y} \bar{F}_t(\kappa y) \right) &\leq \log (e^{\kappa M} \bar{F}_t(\kappa)) = \kappa M + \log \bar{F}_t(\kappa) \\ &= (\log \bar{F}_t(\kappa) + \kappa) + (M - 1)\kappa, \end{aligned}$$

and since $\log \bar{F}_t(\kappa) + \kappa \rightarrow -\infty$ by (5.4), while $(M - 1)\kappa$ is bounded, (5.28) follows.

To prove (5.28) in the regime $\kappa \rightarrow \infty$, we are going to exploit the assumption (2.13). First we fix $\delta > 0$, to be defined later, and set $\bar{n} := \lceil \frac{M-1}{\delta} \rceil$ and $a_n := 1 + n\delta$ for $n = 0, \dots, \bar{n}$, so that $[1, M] \subseteq \bigcup_{n=1}^{\bar{n}} [a_{n-1}, a_n]$. For all $y \in [a_{n-1}, a_n]$ one has, by (2.7),

$$\log \bar{F}_t(\kappa y) \leq \log \bar{F}_t(\kappa a_{n-1}) \sim I_+(a_{n-1}) \log \bar{F}_t(\kappa) \leq a_{n-1} \log \bar{F}_t(\kappa),$$

having used that $I_+(\varrho) \geq \varrho$, by (2.13), hence eventually

$$\log \bar{F}_t(\kappa y) \leq (1 - \delta) a_{n-1} \log \bar{F}_t(\kappa), \quad \forall y \in [a_{n-1}, a_n].$$

Recalling that $a_n = a_{n-1} + \delta$, we can write $a_n \leq (1 - \delta)a_{n-1} + \delta(1 + M)$, because $a_{n-1} \leq M$ by construction, and since $e^{\kappa y} \leq e^{\kappa a_n}$ for $y \in [a_{n-1}, a_n]$, it follows that

$$\begin{aligned} \log \left(\sup_{y \in [1, M]} e^{\kappa y} \bar{F}_t(\kappa y) \right) &\leq \max_{n=1, \dots, \bar{n}} (a_n \kappa + (1 - \delta)a_{n-1} \log \bar{F}_t(\kappa)) \\ &= \max_{n=1, \dots, \bar{n}} ((1 - \delta)a_{n-1} (\log \bar{F}_t(\kappa) + \kappa) + \delta(1 + M)\kappa). \end{aligned}$$

Plainly, the max is attained for $n = 1$, for which $a_{n-1} = a_0 = 1$. Recalling (5.27), we get

$$\log \left(\sup_{y \in [1, M]} e^{\kappa y} \bar{F}_t(\kappa y) \right) \leq (1 - \delta(1 + A + AM)) (\log \bar{F}_t(\kappa) + \kappa).$$

Choosing $\delta := \varepsilon/(1 + A + AM)$, the claim (5.28) is proved.

We are ready to give sharp upper bounds on $c(\kappa, t)$, refining (5.19). For fixed $M \in (0, \infty)$, we write

$$c(\kappa, t) = \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{\kappa < X_t \leq \kappa M\}}] + \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa M\}}], \quad (5.29)$$

and we estimate the first term as follows: by Fubini-Tonelli's theorem and (5.28),

$$\begin{aligned} \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{\kappa < X_t \leq \kappa M\}}] &= \mathbb{E} \left[\left(\int_\kappa^\infty e^x \mathbf{1}_{\{x < X_t\}} dx \right) \mathbf{1}_{\{\kappa < X_t \leq \kappa M\}} \right] \\ &= \int_\kappa^{\kappa M} e^x \mathbb{P}(x < X_t \leq \kappa M) dx \leq \int_\kappa^{\kappa M} e^x \bar{F}_t(x) dx \\ &= \kappa \int_1^M e^{\kappa y} \bar{F}_t(\kappa y) dy \leq \kappa (M - 1) e^{(1 - \varepsilon)(\log \bar{F}_t(\kappa) + \kappa)}. \end{aligned} \quad (5.30)$$

To estimate the second term in (5.29), we start with the cases $\kappa \rightarrow \infty$ and $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$, where we assume that (2.9) holds for some $\eta > 0$, as well as (2.16), hence we can fix $M > 1$ such that $I_+(M) > \frac{1+\eta}{\eta}$. Bounding $(e^{X_t} - e^\kappa) \leq e^{X_t}$, Hölder's inequality yields

$$\mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa M\}}] \leq \mathbb{E}[e^{(1+\eta)X_t}]^{\frac{1}{1+\eta}} \bar{F}_t(\kappa M)^{\frac{\eta}{1+\eta}} = C \bar{F}_t(\kappa M)^{\frac{\eta}{1+\eta}},$$

where $C \in (0, \infty)$ is an absolute constant, by (2.9). Applying relation (2.7) together with $I_+(M) > \frac{1+\eta}{\eta}$ we obtain

$$\frac{\eta}{1 + \eta} \log \bar{F}_t(\kappa M) \sim \frac{\eta}{1 + \eta} I_+(M) \log \bar{F}_t(\kappa) \leq \log \bar{F}_t(\kappa), \quad (5.31)$$

hence eventually

$$\log \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa M\}}] \leq (1 - \varepsilon) \log \bar{F}_t(\kappa) \leq (1 - \varepsilon)(\log \bar{F}_t(\kappa) + \kappa). \quad (5.32)$$

Recalling (5.6) and (5.4), eventually $\kappa(M - 1) \leq e^{-\varepsilon(\log \bar{F}_t(\kappa) + \kappa)}$, hence by (5.30)

$$\log \mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{\kappa < X_t \leq \kappa M\}}] \leq (1 - 2\varepsilon)(\log \bar{F}_t(\kappa) + \kappa). \quad (5.33)$$

Looking back at (5.29), since

$$\log(a + b) \leq \log 2 + \max\{\log a, \log b\}, \quad \forall a, b > 0, \quad (5.34)$$

by (5.32), (5.33) and again (5.4) one has eventually

$$\log c(\kappa, t) \leq \log 2 + (1 - 2\varepsilon)(\log \bar{F}_t(\kappa) + \kappa) \leq (1 - 3\varepsilon)(\log \bar{F}_t(\kappa) + \kappa).$$

Since $\varepsilon > 0$ is arbitrary, this shows that

$$\liminf \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa) + \kappa} \geq 1, \quad (5.35)$$

which together with (5.11) completes the proof of (2.14), if $\kappa \rightarrow \infty$. Since $\log \bar{F}_t(\kappa) + \kappa \sim \log \bar{F}_t(\kappa)$ if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$, by (5.3), we can rewrite (5.35) in this case as

$$\liminf \frac{\log c(\kappa, t)}{\log \bar{F}_t(\kappa)} \geq 1, \quad (5.36)$$

which together with (5.12) completes the proof of (2.17).

It remains to consider the case when $\kappa \rightarrow 0$, where we assume that relation (2.11) holds for some $\eta \in (0, \infty)$, together with (2.16). As before, we fix $M > 1$ such that $I_+(M) > \frac{1+\eta}{\eta}$. Since

$$\mathbb{E}\left[\left(\frac{e^{X_t} - e^\kappa}{\kappa}\right)^{1+\eta} \mathbf{1}_{\{X_t > \kappa\}}\right] \leq \mathbb{E}\left[\left|\frac{e^{X_t} - 1}{\kappa}\right|^{1+\eta}\right] \leq C, \quad (5.37)$$

for some absolute constant $C \in (0, \infty)$, by (2.11), the second term in (5.29) is bounded by

$$\mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa M\}}] \leq \kappa \mathbb{E}\left[\left|\frac{e^{X_t} - e^\kappa}{\kappa}\right|^{1+\eta}\right]^{\frac{1}{1+\eta}} \bar{F}_t(\kappa M)^{\frac{\eta}{1+\eta}} \leq \kappa C \bar{F}_t(\kappa M)^{\frac{\eta}{1+\eta}}. \quad (5.38)$$

In complete analogy with (5.31)-(5.32), we obtain that eventually

$$\log \frac{\mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{X_t > \kappa M\}}]}{\kappa} \leq (1 - \varepsilon) \log \bar{F}_t(\kappa). \quad (5.39)$$

By (5.4), eventually $(M - 1) \leq e^{-\varepsilon(\log \bar{F}_t(\kappa) + \kappa)}$, hence by (5.30)

$$\log \frac{\mathbb{E}[(e^{X_t} - e^\kappa) \mathbf{1}_{\{\kappa < X_t \leq \kappa M\}}]}{\kappa} \leq (1 - 2\varepsilon)(\log \bar{F}_t(\kappa) + \kappa). \quad (5.40)$$

Recalling (5.29) and (5.34), we can finally write

$$\log \frac{c(\kappa, t)}{\kappa} \leq \log 2 + (1 - 2\varepsilon)(\log \bar{F}_t(\kappa) + \kappa) \leq (1 - 3\varepsilon) \log \bar{F}_t(\kappa),$$

because $\kappa \rightarrow 0$ and $\log \bar{F}_t(\kappa) \rightarrow -\infty$. Since $\varepsilon > 0$ is arbitrary, we have proved that

$$\liminf \frac{\log (c(\kappa, t)/\kappa)}{\log \bar{F}_t(\kappa)} \geq 1, \quad (5.41)$$

which together with (5.13) completes the proof of (2.19).

Step 4. Upper bounds on $p(-\kappa, t)$. We are going to prove sharp upper bounds on $p(-\kappa, t)$, that will complete the proof of relations (2.21), (2.23) and (2.26).

By (2.1) we can write

$$p(-\kappa, t) = \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t \leq -\kappa\}}] \leq e^{-\kappa} F_t(-\kappa),$$

therefore

$$\frac{\log p(-\kappa, t)}{\log F_t(-\kappa) - \kappa} \geq 1, \quad (5.42)$$

which together with (5.16) completes the proof of (2.26), if $\kappa \rightarrow \infty$. On the other hand, if $\kappa \rightarrow \bar{\kappa} \in (0, \infty)$, since relation (5.42) implies (recall that $\kappa \geq 0$)

$$\frac{\log p(-\kappa, t)}{\log F_t(-\kappa)} \geq 1, \quad (5.43)$$

in view of (5.17), the proof of (2.23) is completed.

It remains to consider the case $\kappa \rightarrow 0$. If relation (2.11) holds for every $\eta \in (0, \infty)$, we argue in complete analogy with (5.22)-(5.23)-(5.24), getting

$$\liminf \frac{\log (p(-\kappa, t)/\kappa)}{\log F_t(-\kappa)} \geq 1, \quad (5.44)$$

which together with (5.18) completes the proof of (2.26). If, on the other hand, relation (2.11) holds only for some $\eta \in (0, \infty)$, we also assume that condition (2.25) holds, hence we can fix $M > 1$ such that $I_-(M) > \frac{1+\eta}{\eta}$. Let us write

$$p(-\kappa, t) = \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{-\kappa M < X_t \leq -\kappa\}}] + \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t \leq -\kappa M\}}]. \quad (5.45)$$

In analogy with (5.30), for every fixed $\varepsilon > 0$, the first term in the right hand side can be estimated as follows (note that $y \mapsto F_t(-\kappa y)$ is decreasing):

$$\begin{aligned} \mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{-\kappa M < X_t \leq -\kappa\}}] &\leq \int_{-\kappa M}^{-\kappa} e^x F_t(x) dx = \kappa \int_1^M e^{-\kappa y} F_t(-\kappa y) dy \\ &\leq \kappa(M-1) F_t(-\kappa) \leq \kappa e^{(1-\varepsilon) \log F_t(-\kappa)}. \end{aligned}$$

The second term in (5.45) is estimated in complete analogy with (5.37)-(5.38)-(5.39), yielding

$$\log \frac{\mathbb{E}[(e^{-\kappa} - e^{X_t}) \mathbf{1}_{\{X_t \leq -\kappa M\}}]}{\kappa} \leq (1-\varepsilon) \log F_t(-\kappa).$$

Recalling (5.34), we obtain from (5.45)

$$\log \frac{p(-\kappa, t)}{\kappa} \leq \log 2 + (1-\varepsilon) \log F_t(-\kappa) \leq (1-2\varepsilon) \log F_t(-\kappa),$$

and since $\varepsilon > 0$ is arbitrary we have proved that relation (5.44) still holds, which together with (5.18) completes the proof of (2.26), and of the whole Theorem 2.3. \square

5.2. Proof of Theorem 2.7. By Skorokhod's representation theorem, we can build a coupling of the random variables $(X_t)_{t \geq 0}$ and Y such that relation (2.29) holds a.s.. Since the function $z \mapsto z^+$ is continuous, recalling that $\gamma_t \rightarrow 0$, for $\kappa \sim a\gamma_t$ we have a.s.

$$\frac{(e^{X_t} - e^\kappa)^+}{\gamma_t} = \left(\frac{e^{Y\gamma_t(1+o(1))} - 1}{\gamma_t} - \frac{e^{a\gamma_t(1+o(1))} - 1}{\gamma_t} \right)^+ \xrightarrow[t \downarrow 0]{a.s.} (Y - a)^+, \quad (5.46)$$

and analogously for $\kappa \sim -a\gamma_t$

$$\frac{(e^\kappa - e^{X_t})^+}{\gamma_t} \xrightarrow[t \downarrow 0]{a.s.} (-a - Y)^+ = (Y + a)^-. \quad (5.47)$$

Taking the expectation of both sides of these relations, one would obtain precisely (2.34). To justify the interchanging of limit and expectation, we observe that the left hand sides of (5.46) and (5.47) are uniformly integrable, being bounded in $L^{1+\eta}$. In fact

$$\frac{|e^{X_t} - e^\kappa|}{\gamma_t} \leq \frac{|e^{X_t} - 1|}{\gamma_t} + \frac{|e^\kappa - 1|}{\gamma_t},$$

and the second term in the right hand side is uniformly bounded (recall that $\kappa \sim a\gamma_t$ by assumption), while the first term is bounded in $L^{1+\eta}$, by (2.31). \square

APPENDIX A. MISCELLANEA

A.1. About conditions (2.3) and (2.4). Recall from §2.1 that $(X_t)_{t \geq 0}$ denotes the risk-neutral log-price, and assume that $X_t \rightarrow X_0 := 0$ in distribution as $t \rightarrow 0$ (which is automatically satisfied if X has right-continuous paths). For an arbitrary family of values of (κ, t) , with $t > 0$ and $\kappa \geq 0$, we show that condition (2.3) implies (2.4).

Assume first that $t \rightarrow 0$ (with no assumption on κ). Since $\kappa \geq 0$, one has $(e^{X_t} - e^\kappa)^+ \rightarrow (1 - e^\kappa)^+ = 0$ in distribution, hence $c(\kappa, t) \rightarrow 0$ by (2.1) and Fatou's lemma. With analogous arguments, one has $p(-\kappa, t) \rightarrow 0$, hence (2.4) is satisfied.

Next we assume that $\kappa \rightarrow \infty$ and t is bounded, say $t \in (0, T]$ for some fixed $T > 0$. Since $z \mapsto (z - c)^+$ is a convex function and $(e^{X_t})_{t \geq 0}$ is a martingale, the process $((e^{X_t} - e^\kappa)^+)_{t \geq 0}$ is a submartingale and by (2.1) we can write

$$0 \leq c(\kappa, t) \leq \mathbb{E}[(e^{X_T} - e^\kappa)^+] = \mathbb{E}[(e^{X_T} - e^\kappa) \mathbf{1}_{\{X_T > \kappa\}}] \leq \mathbb{E}[e^{X_T} \mathbf{1}_{\{X_T > \kappa\}}].$$

It follows that, if $\kappa \rightarrow +\infty$, then $c(\kappa, t) \rightarrow 0$. With analogous arguments, one shows that $p(-\kappa, t) \rightarrow 0$, hence condition (2.4) holds.

A.2. About Remark 2.8. Let $(S_t)_{t \geq 0}$ be a positive process which solves (2.36). By Ito's formula, the process $X_t := \log S_t$ solves

$$\begin{cases} dX_t = \sqrt{V_t} dW_t - \frac{1}{2} V_t dt \\ X_0 = 0 \end{cases}. \quad (A.1)$$

Assuming $V_t \rightarrow \sigma_0^2$ a.s. as $t \rightarrow 0$, we want to show that

$$\frac{X_t}{\sqrt{t}} \xrightarrow[t \rightarrow 0]{d} Y \sim N(0, \sigma_0^2). \quad (A.2)$$

Let us define

$$J_t := \frac{1}{2\sqrt{t}} \int_0^t V_s ds, \quad I_t := \frac{X_t}{\sqrt{t}} + J_t - \sigma_0 \frac{W_t}{\sqrt{t}} = \int_0^t \frac{\sqrt{V_s} - \sigma_0}{\sqrt{t}} dW_s.$$

By $V_t \rightarrow \sigma_0^2$ a.s. it follows that $J_t \sim \frac{\sqrt{t}}{2} \sigma_0 \rightarrow 0$ a.s., by the fundamental theorem of calculus. Moreover, since $\sqrt{V_t} \rightarrow \sigma_0$ a.s.,

$$\langle I \rangle_t := \int_0^t \frac{|\sqrt{V_s} - \sigma_0|^2}{t} ds \leq \sup_{0 \leq s \leq t} |\sqrt{V_s} - \sigma_0|^2 \xrightarrow[t \rightarrow 0]{a.s.} 0. \quad (A.3)$$

We now use the inequality $P(|I_t| > \varepsilon) \leq \frac{\delta}{\varepsilon^2} + P(\langle I \rangle_t > \delta)$, cf. [KS88, Problem 5.25]. Sending first $t \rightarrow 0$ for fixed $\delta > 0$, and then $\delta \rightarrow 0$, we see by (A.3) that $I_t \rightarrow 0$ in probability as $t \rightarrow 0$. Since $\sigma_0 W_t / \sqrt{t} \rightarrow Y \sim N(0, \sigma_0^2)$ in distribution as $t \rightarrow 0$,[†] by Slutsky's theorem

$$\frac{X_t}{\sqrt{t}} = I_t - J_t + \sigma_0 \frac{W_t}{\sqrt{t}} \xrightarrow[t \rightarrow 0]{d} 0 + Y = Y \sim N(0, \sigma_0^2),$$

hence relation (A.2) holds.

Next we show that, plugging $Y \sim N(0, \sigma_0^2)$ into (2.35), we obtain $C_{\pm}(a) = \sigma_0$ for any $a \geq 0$. Since Y has a symmetric law, it suffices to focus on $C_+(a)$. Then

$$\begin{aligned} E[(Y - a)^+] &= \sigma_0 E \left[\left(N(0, 1) - \frac{a}{\sigma_0} \right)^+ \right] = \sigma_0 \left[\int_{\frac{a}{\sigma_0}}^{\infty} x \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} dx - \frac{a}{\sigma_0} \int_{\frac{a}{\sigma_0}}^{\infty} \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} dx \right] \\ &= \sigma_0 \left(\phi \left(\frac{a}{\sigma_0} \right) - \frac{a}{\sigma_0} \Phi \left(-\frac{a}{\sigma_0} \right) \right) = a D \left(\frac{a}{\sigma_0} \right), \end{aligned} \quad (\text{A.4})$$

where we used the density ϕ and distribution function Φ of a $N(0, 1)$ random variable, cf. (4.1), and the definition (2.32) of D . Looking back at (2.35), we obtain $C_+(a) = \sigma_0$. \square

A.3. Proof of Lemma 3.3. We start with some estimates. It follows by (3.20) that

$$X_t \xrightarrow{d} \sigma W_t + \mu t + \sum_{i=1}^{N_t} Y_i$$

with $Y_i \sim N(\alpha, \delta^2)$ and $N_t \sim Pois(\lambda t)$ (and we agree that the sum equals 0 in case $N_t = 0$). By Chernoff's bound[‡] $P(N_t > M) \leq \left(\frac{e\lambda t}{M}\right)^M$, hence

$$P(X_t > \kappa) = e^{-\lambda t} \sum_{n=0}^M P(N(\mu t + n\alpha, \sigma^2 t + n\delta^2) > \kappa) \frac{(\lambda t)^n}{n!} + \mathcal{O} \left(\left(\frac{e\lambda t}{M} \right)^M \right), \quad (\text{A.5})$$

where $N(a, b)$ denotes a Gaussian random variable with mean a and variance b . We recall the standard estimate $\log P(N(0, 1) > x) \sim -\frac{1}{2}x^2$ as $x \rightarrow \infty$. Then we can write:

If t is bounded from above (e.g. $t \rightarrow \bar{t} \in [0, \infty)$) and $\kappa \rightarrow \infty$,

$$\log P(N(\mu t + n\alpha, \sigma^2 t + n\delta^2) > \kappa) \sim -\frac{\kappa^2}{2(\sigma^2 t + n\delta^2)}. \quad (\text{A.6})$$

In particular, we get from (A.5) that, for fixed $M \in \mathbb{N}$,

$$\begin{aligned} P(X_t > \kappa) &\sim e^{-\frac{\kappa^2}{2\sigma^2 t}(1+o(1))} + \sum_{n=1}^M e^{-\frac{\kappa^2}{2(\sigma^2 t + n\delta^2)}(1+o(1))} \frac{(\lambda t)^n}{n!} + \mathcal{O} \left(\left(\frac{e\lambda t}{M} \right)^M \right) \\ &\leq e^{-\frac{\kappa^2}{2\sigma^2 t}(1+o(1))} + M \max_{n=1, \dots, M} e^{-\left(\frac{\kappa^2}{2n\delta^2} + n \log \frac{1}{\lambda t} + \log n! \right)(1+o(1))} + \mathcal{O} \left(\left(\frac{e\lambda t}{M} \right)^M \right). \end{aligned} \quad (\text{A.7})$$

For a lower bound, restricting the sum in (A.5) to a single value $n \in \mathbb{N}$, we get

$$P(X_t > \kappa) \geq e^{-\left(\frac{\kappa^2}{2(\sigma^2 t + n\delta^2)} + n \log \frac{1}{\lambda t} + \log n! \right)(1+o(1))}. \quad (\text{A.8})$$

[†]In fact, the distribution of $\sigma_0 W_t / \sqrt{t}$ is $N(0, \sigma_0^2)$ for all $t > 0$.

[‡]Just apply Markov's inequality $P(N_t > M) \leq e^{-M\alpha} E[e^{\alpha N_t}] = e^{-M\alpha + \lambda t(e^\alpha - 1)}$ and optimize over $\alpha \geq 0$.

We now prove relation (3.24). We fix a family of (κ, t) with $t \rightarrow 0$ and $\kappa \sim a \kappa_2(t)$ for some $a \in (0, \infty)$. To get an upper bound, we drop the term $\log n!$ in (A.7) (since $e^{-\log n!} \leq 1$) and plug $\kappa \sim a \sqrt{\log \frac{1}{t}}$, getting

$$P(X_t > \kappa) \leq t^{\frac{a^2}{2\sigma^2 t}(1+o(1))} + M \max_{n=1, \dots, M} t^{\left(\frac{a^2}{2n\delta^2} + n\right)(1+o(1))} + \mathcal{O}\left(\left(\frac{e\lambda t}{M}\right)^M\right). \quad (\text{A.9})$$

Let us denote by $\bar{n}_a \in \mathbb{N}$ the value of $n \in \mathbb{N}$ for which the minimum in the definition (3.22) of $f(a)$ is attained. Choosing $M \in \mathbb{N}$ large enough, so that $M \geq \bar{n}_a$, the middle term in (A.9) is $t^{f(a)(1+o(1))}$ and is the dominating one, provided $M > e\lambda$ and $M > f(a)$, so that the third term is $\ll t^{f(a)}$. For an analogous lower bound, we apply (A.8) with $n = \bar{n}_a$: since $\sigma^2 t + n\delta^2 \sim n\delta^2$ (recall that $t \rightarrow 0$), we get

$$P(X_t > \kappa) \geq e^{-\log \bar{n}_a!} t^{f(a)(1+o(1))} = (\text{const.}) t^{f(a)(1+o(1))}.$$

We have thus proved relation (3.24).

It remains to prove relation (3.25). We fix a family of (κ, t) such that either $t \rightarrow 0$ and $\kappa \gg \kappa_2(t)$, or $t \rightarrow \bar{t} \in (0, \infty)$ and $\kappa \rightarrow \infty$. Since $n! \geq (n/e)^n$,

$$\frac{\kappa^2}{2n\delta^2} + n \log \frac{1}{\lambda t} + \log n! \geq \frac{\kappa^2}{2n\delta^2} + n \log \frac{n}{e\lambda t} \geq \inf_{x \geq 0} \left\{ \frac{\kappa^2}{2\delta^2 x} + x \log \frac{x}{e\lambda t} \right\}.$$

By direct computation, the infimum is attained at

$$\bar{x} \sim \frac{\kappa}{\delta \sqrt{2 \log \frac{\kappa}{t}}}, \quad (\text{A.10})$$

which yields

$$\frac{\kappa^2}{2n\delta^2} + n \log \frac{1}{\lambda t} + \log n! \geq \frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}} (1 + o(1)).$$

We now choose $M = \lfloor 3\bar{x} \rfloor$ in (A.7), so that

$$\begin{aligned} P(X_t > \kappa) &\leq e^{-\frac{\kappa^2}{2\sigma^2 t}(1+o(1))} + 3\bar{x} e^{-\frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}} (1+o(1))} + \mathcal{O}\left(\left(\frac{\lambda t}{\bar{x}}\right)^{3\bar{x}}\right) \\ &\leq e^{-\frac{\kappa^2}{2\sigma^2 t}(1+o(1))} + e^{-\frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}} (1+o(1))} + \mathcal{O}\left(e^{-3\bar{x} \log \frac{\bar{x}}{\lambda t}}\right), \end{aligned} \quad (\text{A.11})$$

where we have absorbed $3\bar{x}$ inside the $o(1)$ term in the exponential, because $\log(3\bar{x}) = o(\kappa) = o(\kappa \sqrt{\log \frac{\kappa}{t}})$ by (A.10) (recall that $\kappa \rightarrow \infty$). The dominant contribution to (A.11) is given by the middle term (note that $3\bar{x} \log \frac{\bar{x}}{\lambda t} \sim \frac{3}{2} \frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}}$, always by (A.10)). For a corresponding lower bound, we apply (A.8) with $n = \lfloor \bar{x} \rfloor$: since $\log n! \sim n \log(n/e)$ and $\sigma^2 t + \lfloor \bar{x} \rfloor \delta^2 \sim \lfloor \bar{x} \rfloor \delta^2$ (because $\bar{x} \rightarrow \infty$), we get

$$\begin{aligned} P(X_t > \kappa) &\geq e^{-\left(\frac{\kappa^2}{2\delta^2 \lfloor \bar{x} \rfloor} + \lfloor \bar{x} \rfloor \log \frac{1}{\lambda t} + \log \lfloor \bar{x} \rfloor!\right)(1+o(1))} = e^{-\left(\frac{\kappa^2}{2\delta^2 \lfloor \bar{x} \rfloor} + \lfloor \bar{x} \rfloor \log \frac{\lfloor \bar{x} \rfloor}{e\lambda t}\right)(1+o(1))} \\ &= e^{-\frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}} (1+o(1))}. \end{aligned}$$

We have thus shown that

$$\log P(X_t > \kappa) \sim -\frac{\kappa}{\delta} \sqrt{2 \log \frac{\kappa}{t}},$$

completing the proof of relation (3.25) and of Lemma 3.3. \square

A.4. Proof of Proposition 4.2. Let us first prove (4.11) and (4.12). Since $\phi(d_2)e^\kappa = \phi(d_1)$, cf. (4.1) and (4.8), recalling (4.2) we can rewrite the Black&Scholes formula (4.7) as follows:

$$C_{BS}(\kappa, v) = \phi(d_1)(U(-d_1) - U(-d_2)) = \phi(d_1)(U(-d_1) - U(-d_1 + v)). \quad (\text{A.12})$$

If $d_1 \rightarrow -\infty$, applying (4.3) we get

$$U(-d_1) - U(-d_1 + v) = - \int_{-d_1}^{-d_1+v} U'(z) dz \sim \int_{-d_1}^{-d_1+v} \frac{1}{z^2} dz = \frac{v}{-d_1(-d_1+v)},$$

and (4.11) is proved. Next we assume that $v \rightarrow 0$. By convexity of $U(\cdot)$ (cf. Lemma 4.1),

$$-U'(-d_1 + v) \leq \frac{U(-d_1) - U(-d_1 + v)}{v} \leq -U'(-d_1),$$

hence to prove (4.12) it suffices to show that $U'(-d_1 + v) \sim U'(-d_1)$. To this purpose, by a subsequence argument, we may assume that $d_1 \rightarrow \bar{d}_1 \in \mathbb{R} \cup \{\pm\infty\}$. Since $d_1 \leq \frac{v}{2}$ for $\kappa \geq 0$, when $v \rightarrow 0$ necessarily $\bar{d}_1 \in [-\infty, 0]$. If $\bar{d}_1 = -\infty$, i.e. $-d_1 \rightarrow +\infty$, then $-d_1 + v \sim -d_1 \rightarrow +\infty$ and $U'(-d_1 + v) \sim U'(-d_1)$ follows by (4.3). On the other hand, if $\bar{d}_1 \in (-\infty, 0]$ then both $U'(-d_1)$ and $U'(-d_1 + v)$ converge to $U'(-\bar{d}_1) \neq 0$, by continuity of U' , hence $U'(-d_1)/U'(-d_1 + v) \rightarrow 1$, i.e. $U'(-d_1 + v) \sim U'(-d_1)$ as requested.

Let us now prove (4.10). Assume that $\min\{d_1, \log v\} \rightarrow -\infty$, and note that for every subsequence we can extract a sub-subsequence along which either $d_1 \rightarrow -\infty$ or $v \rightarrow 0$. We can then apply (4.11) and (4.12) to show that $C_{BS}(\kappa, v) \rightarrow 0$:

- if $d_1 \rightarrow -\infty$, the right hand side of (4.11) is bounded from above by $\phi(d_1)/(-d_1) \rightarrow 0$;
- If $\kappa \geq 0$ and $v \rightarrow 0$, then $d_1 \leq \frac{v}{2} \rightarrow 0$ and consequently $\phi(d_1)U'(-d_1)$ is uniformly bounded from above, hence the right hand side of (4.12) vanishes (since $v \rightarrow 0$).

Finally, we assume that $\min\{d_1, \log v\} \not\rightarrow -\infty$ and show that $C_{BS}(\kappa, v) \not\rightarrow 0$. Extracting a subsequence, we have $\min\{d_1, \log v\} \geq -M$ for some fixed $M \in (0, \infty)$, i.e. both $v \geq \varepsilon := e^{-M} > 0$ and $d_1 \geq -M$, and we may assume that $v \rightarrow \bar{v} \in [\varepsilon, +\infty]$ and $d_1 \rightarrow \bar{d}_1 \in [-M, +\infty]$. Consider first the case $\bar{v} = +\infty$, i.e. $v \rightarrow +\infty$: by (4.8) one has $-d_1 + v = -d_2 \geq \frac{v}{2} \rightarrow +\infty$, hence $\phi(d_1)U(-d_1 + v) \rightarrow 0$ (because ϕ is bounded), and recalling (4.2) relation (A.12) yields

$$C_{BS}(\kappa, v) = \Phi(d_1) - \phi(d_1)U(-d_1 + v) \rightarrow \Phi(\bar{d}_1) > 0.$$

Next consider the case $\bar{v} < +\infty$: since $d_1 \leq \frac{v}{2}$, we have $\bar{d}_1 \leq \frac{\bar{v}}{2}$ and again by (A.12) we obtain $C_{BS}(\kappa, v) \rightarrow \phi(\bar{d}_1)(U(-\bar{d}_1) - U(-\bar{d}_1 + \bar{v})) > 0$. In both cases, $C_{BS}(\kappa, v) \not\rightarrow 0$. \square

ACKNOWLEDGMENTS

We thank Fabio Bellini, Stefan Gerhold and Carlo Sgarra for fruitful discussions.

REFERENCES

- [AP07] L. B. G. Andersen and V. V. Piterbarg (2007), *Moment explosions in stochastic volatility models*, Finance Stoch. **11**, pp. 29–50.
- [ACDP12] A. Andreoli, F. Caravenna, P. Dai Pra and G. Posta (2012), *Scaling and multiscaling in financial series: a simple model*, Adv. in Appl. Probab. **44**, pp. 1018–1051.
- [BF08] S. Benaim and P. Friz (2008), *Smile asymptotics II: models with known moment generating function*, J. Appl. Probab. **45**, pp. 16–32.
- [BF09] S. Benaim and P. Friz (2009), *Regular variation and smile asymptotics*, Math. Finance **19**, pp. 1–12.

- [BGT89] N. H. Bingham, C. H. Goldie and J. L. Teugels (1989), *Regular variation*, Cambridge University Press.
- [CC15] F. Caravenna and J. Corbetta (2015), *The asymptotic smile of a multiscaling stochastic volatility model*, preprint, arXiv.org 1501.03387 [math.PR].
- [CW04] P. Carr and L. Wu (2004), *The Finite Moment Log Stable Process and Option Pricing*, J. Finance **58**, pp. 753–778.
- [DZ98] A. Dembo and O. Zeitouni (1998), *Large deviations techniques and applications*, Second Edition, Springer–Verlag, New York.
- [FH09] J. E. Figueroa-López, and C. Houdré (2009), *Small-time expansions for the transition distributions of Lévy processes* Stochastic Processes and their applications **119.11**, pp. 3862–3889.
- [FF12] J. E. Figueroa-López, and M. Forde (2012), *The Small-Maturity Smile for Exponential Lévy Models*, SIAM J. Financial Math. **3**, pp. 33–65.
- [FGY14] P. Friz, S. Gerhold and M. Yor (2014), *How to make Dupire’s local volatility work with jumps* Quantitative Finance, DOI: 10.1080/14697688.2013.874622
- [FJ09] M. Forde and A. Jacquier (2009), *Small-time asymptotics for implied volatility under the Heston model*, Int J. Theor. Appl. Finance **12**, pp. 861–876
- [FK09] P. Friz and M. Keller-Ressel (2009), *Moment Explosions in Stochastic Volatility Models*, Contribution to Encyclopedia of Quantitative Finance.
- [GL14] K. Gao, R. Lee (2014), *Asymptotics of implied volatility to arbitrary order*, Finance Stoch. **18**, pp. 349–392.
- [G11] J. Gatheral (2011), *The Volatility Surface: A Practitioner’s Guide*, John Wiley & Sons.
- [GMZ14] S. Gerhold, J. F. Morgenbesser and A. Zrunek (2014), *Refined wing asymptotics for the Merton and Kou jump diffusion models*, to appear on Banach Center Publications, arXiv.org: 1401.1954.
- [G10] A. Gulisashvili (2010), *Asymptotic formulas with error estimates for call pricing functions and the implied volatility at extreme strikes*, SIAM J. Financial Math. **1.1**, pp. 609–641.
- [H93] S. L. Heston (1993), *A closed-form solution for options with stochastic volatility and applications to bond and currency options*, Rev. Financ. Stud. **6**, pp. 327–343.
- [JKM13] A. Jacquier, M. Keller-Ressel and A. Mijatović (2013), *Large deviations and stochastic volatility with jumps: asymptotic implied volatility for affine models*, Stochastics **85**, pp. 321–345.
- [KS88] I. Karatzas, S. E. Shreve, *Brownian Motion and Stochastic Calculus*, Springer (1988).
- [L04] R. Lee (2004), *The moment formula for implied volatility at extreme strikes*, Math. Finance **14**, pp. 469–480.
- [M76] R. C. Merton (1976), *Option pricing when underlying stock returns are discontinuous*, J. Financial Economics **3**, pp. 125–144.
- [MT12] A. Mijatović and P. Tankov (2012), *A new look at short-term implied volatility in asset price models with jumps*, to appear on Math. Finance, arXiv:1207.0843.
- [PP15] A. Pascucci and S. Pagliarani (2015), *The parabolic Taylor formula of the implied volatility*, preprint, arXiv.org: 1510.06084 [math.PR]
- [P01] I. Pinelis (2002), *Monotonicity Properties of the Relative Error of a Padé Approximation for Mills’ Ratio*, J. Inequal. Pure Appl. Math. **3**, article 20.
- [RR09] M. Roper and M. Rutkowski (2009), *On the relationship between the call price surface and the implied volatility surface close to expiry*, Int. J. Theor. Appl. Finance **12**, pp. 427–441.
- [S54] L. R. Shenton (1954), *Inequalities for the normal integral including a new continued fraction*, Biometrika **41**, pp. 177–189.
- [T09] M. Tehranchi (2009), *Asymptotics of implied volatility far from maturity*, J. Appl. Probab. **46**, pp. 629–650.

DIPARTIMENTO DI MATEMATICA E APPLICAZIONI, UNIVERSITÀ DEGLI STUDI DI MILANO-BICOCCA, VIA COZZI 55, I-20125 MILANO, ITALY

E-mail address: francesco.caravenna@unimib.it

ÉCOLE DES PONTS - PARIS TECH, CERMICS, 6 ET 8 AVENUE BLAISE PASCAL, 77420 CHAMPS SUR MARNE, FRANCE

E-mail address: jacopo.corbetta@enpc.fr